

Claims reserving with R: ChainLadder-0.2.1 Package Vignette

Alessandro Carrato, Fabio Concina, Markus Gesmann, Dan Murphy,
Mario Wüthrich and Wayne Zhang

July 11, 2015

Abstract

The `ChainLadder` package provides various statistical methods which are typically used for the estimation of outstanding claims reserves in general insurance, including those to estimate the claims development results as required under Solvency II.

Contents

1	Introduction	4
1.1	Claims reserving in insurance	4
2	The ChainLadder package	5
2.1	Motivation	5
2.2	Brief package overview	5
2.3	Installation	6
3	Using the ChainLadder package	6
3.1	Working with triangles	6
3.1.1	Plotting triangles	8
3.1.2	Transforming triangles between cumulative and incremental representation	9
3.1.3	Importing triangles from external data sources	10
4	Chain-ladder methods	14
4.1	Basic idea	14
4.2	Mack chain-ladder	18
4.3	Munich chain-ladder	22
4.4	Bootstrap chain-ladder	24
4.5	Multivariate chain-ladder	27
4.6	The "triangles" class	28
4.7	Separate chain ladder ignoring correlations	29
4.8	Multivariate chain ladder using seemingly unrelated regressions	31
4.9	Other residual covariance estimation methods	32
4.10	Model with intercepts	35
4.11	Joint modelling of the paid and incurred losses	37
5	Clark's methods	38
5.1	Clark's LDF method	39
5.2	Clark's Cap Cod method	41

6 Generalised linear model methods	42
7 Paid-incurred chain model	47
7.1 Model assumptions	47
7.2 Parameter estimation	47
8 One year claims development result	49
8.1 CDR functions	50
9 Model Validation with tweedieReserve	52
10 Using ChainLadder with RExcel and SWord	54
11 Further resources	54
11.1 Other insurance related R packages	55
11.2 Presentations	55
11.3 Further reading	56
12 Training and consultancy	56
References	59

1 Introduction

1.1 Claims reserving in insurance

The insurance industry, unlike other industries, does not sell products as such but promises. An insurance policy is a promise by the insurer to the policyholder to pay for future claims for an upfront received premium.

As a result insurers don't know the upfront cost for their service, but rely on historical data analysis and judgement to predict a sustainable price for their offering. In General Insurance (or Non-Life Insurance, e.g. motor, property and casualty insurance) most policies run for a period of 12 months. However, the claims payment process can take years or even decades. Therefore often not even the delivery date of their product is known to insurers.

In particular losses arising from casualty insurance can take a long time to settle and even when the claims are acknowledged it may take time to establish the extent of the claims settlement cost. Claims can take years to materialize. A complex and costly example are the claims from asbestos liabilities, particularly those in connection with mesothelioma and lung damage arising from prolonged exposure to asbestos. A research report by a working party of the Institute and Faculty of Actuaries estimated that the un-discounted cost of UK mesothelioma-related claims to the UK Insurance Market for the period 2009 to 2050 could be around £10bn, see [GBB⁺09]. The cost for asbestos related claims in the US for the worldwide insurance industry was estimate to be around \$120bn in 2002, see [Mic02].

Thus, it should come as no surprise that the biggest item on the liabilities side of an insurer's balance sheet is often the provision or reserves for future claims payments. Those reserves can be broken down in case reserves (or outstanding claims), which are losses already reported to the insurance company and losses that are incurred but not reported (IBNR) yet.

Historically, reserving was based on deterministic calculations with pen and paper, combined with expert judgement. Since the 1980's, with the arrival of personal computer, spreadsheet software became very popular for reserving. Spreadsheets not only reduced the calculation time, but allowed actuaries to test different scenarios and the sensitivity of their forecasts.

As the computer became more powerful, ideas of more sophisticated models started to evolve. Changes in regulatory requirements, e.g. Solvency II¹ in Europe, have fostered further research and promoted the use of stochastic and statistical techniques. In particular, for many countries extreme percentiles of reserve deterioration over a fixed time period have to be estimated for the purpose of capital setting.

Over the years several methods and models have been developed to estimate both the level and variability of reserves for insurance claims, see [Sch11] or [PR02] for an overview.

¹See http://ec.europa.eu/internal_market/insurance/solvency/index_en.htm

In practice the Mack chain-ladder and bootstrap chain-ladder models are used by many actuaries along with stress testing / scenario analysis and expert judgement to estimate ranges of reasonable outcomes, see the surveys of UK actuaries in 2002, [LFK⁺02], and across the Lloyd's market in 2012, [Orr12].

2 The ChainLadder package

2.1 Motivation

The ChainLadder [GMZ⁺15] package provides various statistical methods which are typically used for the estimation of outstanding claims reserves in general insurance. The package started out of presentations given by Markus Gesmann at the Stochastic Reserving Seminar at the Institute of Actuaries in 2007 and 2008, followed by talks at Casualty Actuarial Society (CAS) meetings joined by Dan Murphy in 2008 and Wayne Zhang in 2010.

Implementing reserving methods in R has several advantages. R provides:

- a rich language for statistical modelling and data manipulations allowing fast prototyping
- a very active user base, which publishes many extensions
- many interfaces to data bases and other applications, such as MS Excel
- an established framework for End User Computing, including documentation, testing and workflows with version control systems
- code written in plain text files, allowing effective knowledge transfer
- an effective way to collaborate over the internet
- built in functions to create reproducible research reports²
- in combination with other tools such as L^AT_EX and Sweave or Markdown easy to set up automated reporting facilities
- access to academic research, which is often first implemented in R

2.2 Brief package overview

This vignette will give the reader a brief overview of the functionality of the ChainLadder package. The functions are discussed and explained in more detail in the respective help files and examples, see also [Ges14].

A set of demos is shipped with the packages and the list of demos is available via:

²For an example see the project: Formatted Actuarial Vignettes in R, <http://www.favir.net/>

```
R> demo(package="ChainLadder")
```

and can be executed via

```
R> library(ChainLadder)
R> demo("demo name")
```

For more information and examples see the project web site: <https://github.com/mages/ChainLadder>

2.3 Installation

You can install ChainLadder in the usual way from CRAN, e.g.:

```
R> install.packages('ChainLadder')
```

For more details about installing packages see [Tea12b]. The installation was successful if the command `library(ChainLadder)` gives you the following message:

```
R> library(ChainLadder)
```

Welcome to ChainLadder version 0.2.1

Type `vignette('ChainLadder', package='ChainLadder')` to access the overall package documentation.

See `demo(package='ChainLadder')` for a list of demos.

More information is available on the ChainLadder project web-site: <https://github.com/mages/ChainLadder>

To suppress this message use:
`suppressPackageStartupMessages(library(ChainLadder))`

3 Using the ChainLadder package

3.1 Working with triangles

Historical insurance data is often presented in form of a triangle structure, showing the development of claims over time for each exposure (origin) period. An origin period could be the year the policy was written or earned, or the loss occurrence period. Of course the origin period doesn't have to be yearly, e.g. quarterly or

monthly origin periods are also often used. The development period of an origin period is also called age or lag. Data on the diagonals present payments in the same calendar period. Note, data of individual policies is usually aggregated to homogeneous lines of business, division levels or perils.

Most reserving methods of the ChainLadder package expect triangles as input data sets with development periods along the columns and the origin period in rows. The package comes with several example triangles. The following R command will list them all:

```
R> require(ChainLadder)
R> data(package="ChainLadder")
```

Let's look at one example triangle more closely. The following triangle shows data from the Reinsurance Association of America (RAA):

```
R> ## Sample triangle
R> RAA
```

	dev									
origin	1	2	3	4	5	6	7	8	9	10
1981	5012	8269	10907	11805	13539	16181	18009	18608	18662	18834
1982	106	4285	5396	10666	13782	15599	15496	16169	16704	NA
1983	3410	8992	13873	16141	18735	22214	22863	23466	NA	NA
1984	5655	11555	15766	21266	23425	26083	27067	NA	NA	NA
1985	1092	9565	15836	22169	25955	26180	NA	NA	NA	NA
1986	1513	6445	11702	12935	15852	NA	NA	NA	NA	NA
1987	557	4020	10946	12314	NA	NA	NA	NA	NA	NA
1988	1351	6947	13112	NA	NA	NA	NA	NA	NA	NA
1989	3133	5395	NA	NA	NA	NA	NA	NA	NA	NA
1990	2063	NA	NA	NA	NA	NA	NA	NA	NA	NA

This triangle shows the known values of loss from each origin year and of annual evaluations thereafter. For example, the known values of loss originating from the 1988 exposure period are 1351, 6947, and 13112 as of year ends 1988, 1989, and 1990, respectively. The *latest diagonal* – i.e., the vector 18834, 16704, ... 2063 from the upper right to the lower left – shows the most recent evaluation available. The column headings – 1, 2, ..., 10 – hold the *ages* (in years) of the observations in the column relative to the beginning of the exposure period. For example, for the 1988 origin year, the age of the 1351 value, evaluated as of 1988-12-31, is three years.

The objective of a reserving exercise is to forecast the future claims development in the bottom right corner of the triangle and potential further developments beyond development age 10. Eventually all claims for a given origin period will be settled, but it is not always obvious to judge how many years or even decades it will take.

We speak of long and short tail business depending on the time it takes to pay all claims.

3.1.1 Plotting triangles

The first thing you often want to do is to plot the data to get an overview. For a data set of class `triangle` the `ChainLadder` package provides default plotting methods to give a graphical overview of the data:

```
R> plot(RAA)
```

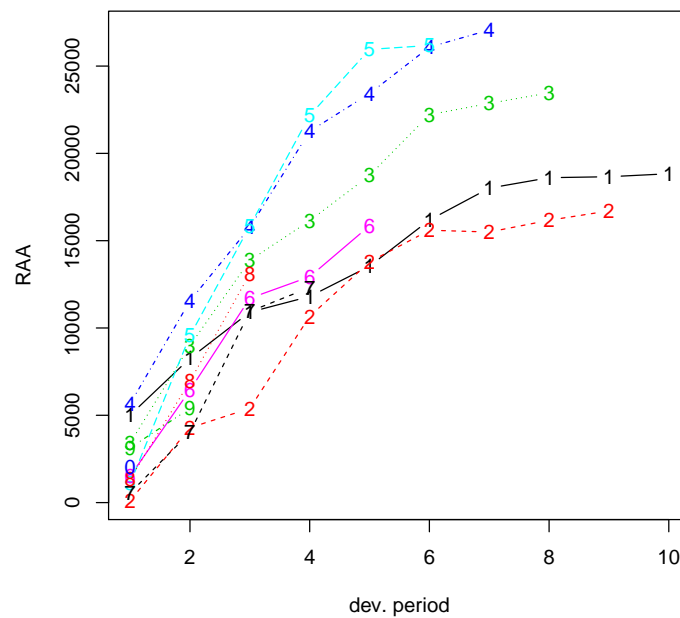


Figure 1: Claims development chart of the RAA triangle, with one line per origin period. Output of `plot(RAA)`

Setting the argument `lattice=TRUE` will produce individual plots for each origin period³, see Figure 2.

```
R> plot(RAA, lattice=TRUE)
```

³`ChainLadder` uses the `lattice` package for plotting the development of the origin years in separate panels.

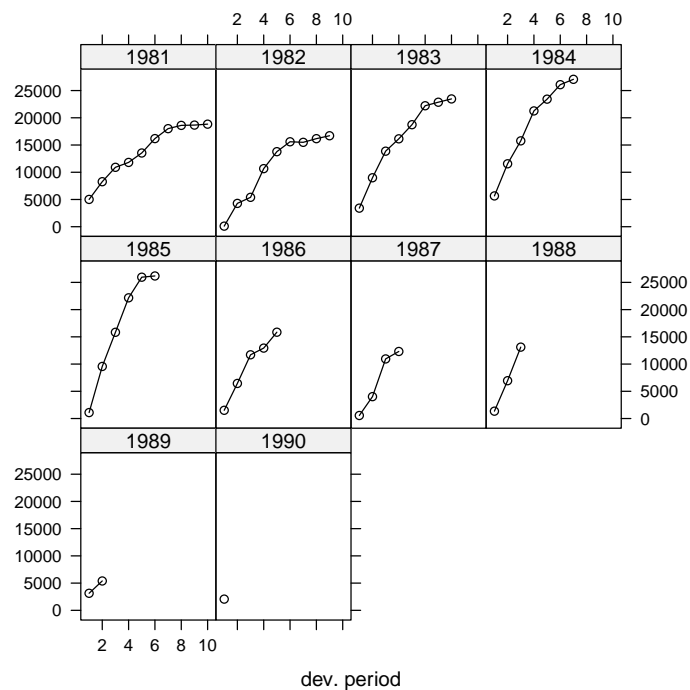


Figure 2: Claims development chart of the RAA triangle, with individual panels for each origin period. Output of `plot(RAA, lattice=TRUE)`

You will notice from the plots in Figures 1 and 2 that the triangle RAA presents claims developments for the origin years 1981 to 1990 in a cumulative form. For more information on the triangle plotting functions see the help pages of `plot.triangle`, e.g. via

```
R> ?plot.triangle
```

3.1.2 Transforming triangles between cumulative and incremental representation

The ChainLadder packages comes with two helper functions, `cum2incr` and `incr2cum` to transform cumulative triangles into incremental triangles and vice versa:

```
R> raa.inc <- cum2incr(RAA)
R> ## Show first origin period and its incremental development
R> raa.inc[1,]
```

1	2	3	4	5	6	7	8	9	10
5012	3257	2638	898	1734	2642	1828	599	54	172

```
R> raa.cum <- incr2cum(raa.inc)
R> ## Show first origin period and its cumulative development
R> raa.cum[1,]
```

1	2	3	4	5	6	7	8	9	10
5012	8269	10907	11805	13539	16181	18009	18608	18662	18834

3.1.3 Importing triangles from external data sources

In most cases you want to analyse your own data, usually stored in data bases or spreadsheets.

Importing a triangle from a spreadsheet There are many ways to import the data from a spreadsheet. A quick and dirty solution is using a CSV-file.

Open a new workbook and copy your triangle into cell A1, with the first column being the accident or origin period and the first row describing the development period or age.

Ensure the triangle has no formatting, such a commas to separate thousands, as those cells will be saved as characters.

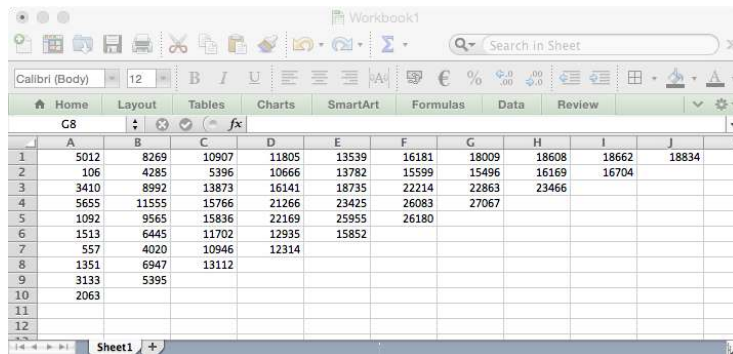


Figure 3: Screen shot of a triangle in a spreadsheet software.

Now open R and go through the following commands:

```
R> myCSVfile <- "path/to/folder/with/triangle.csv"
R> ## Use the R command:
R> # myCSVfile <- file.choose()
R> ## to select the file interactively
```

```
R> tri <- read.csv(file=myCSVfile, header = FALSE)
R> ## Use read.csv2 if semicolons are used as a separator likely
R> ## to be the case if you are in continental Europe
R> library(ChainLadder)
R> ## Convert to triangle
R> tri <- as.triangle(as.matrix(tri))
R> # Job done.
```

Copying and pasting from a spreadsheet Small data sets can be transferred to R backwards and forwards via the clipboard under MS Windows.

Select a data set in the spreadsheet and copy it into the clipboard, then go to R and type:

```
R> tri <- read.table(file="clipboard", sep="\t", na.strings="")
```

Reading data from a data base R makes it easy to access data using SQL statements, e.g. via an ODBC connection⁴, for more details see [Tea12a]. The ChainLadder packages includes a demo to showcase how data can be imported from a MS Access data base, see:

```
R> demo(DatabaseExamples)
```

In this section we use data stored in a CSV-file⁵ to demonstrate some typical operations you will want to carry out with data stored in data bases. CSV stands for comma separated values, stored in a text file. Note many European countries use a comma as decimal point and a semicolon as field separator, see also the help file to read.csv2. In most cases your triangles will be stored in tables and not in a classical triangle shape. The ChainLadder package contains a CSV-file with sample data in a long table format. We read the data into R's memory with the read.csv command and look at the first couple of rows and summarise it:

```
R> filename <- file.path(system.file("Database",
                                     package="ChainLadder"),
                          "TestData.csv")
R> myData <- read.csv(filename)
R> head(myData)
```

	origin	dev	value	lob
1	1977	1	153638	ABC
2	1978	1	178536	ABC

⁴See the [RODBC](#) and [DBI](#) packages

⁵Please ensure that your CSV-file is free from formatting, e.g. characters to separate units of thousands, as those columns will be read as characters or factors rather than numerical values.

```

3  1979  1 210172 ABC
4  1980  1 211448 ABC
5  1981  1 219810 ABC
6  1982  1 205654 ABC

```

```
R> summary(myData)
```

	origin	dev	value		lob
Min. :	1	Min. : 1.00	Min. : -17657	AutoLiab	:105
1st Qu.:	3	1st Qu.: 2.00	1st Qu.: 10324	GeneralLiab	:105
Median :	6	Median : 4.00	Median : 72468	M3IR5	:105
Mean :	642	Mean : 4.61	Mean : 176632	ABC	: 66
3rd Qu.:	1979	3rd Qu.: 7.00	3rd Qu.: 197716	CommercialAutoPaid:	55
Max. :	1991	Max. :14.00	Max. :3258646	GenIns	: 55
				(Other)	:210

Let's focus on one subset of the data. We select the RAA data again:

```
R> raa <- subset(myData, lob %in% "RAA")
R> head(raa)
```

```

      origin dev value lob
67  1981    1  5012 RAA
68  1982    1   106 RAA
69  1983    1  3410 RAA
70  1984    1  5655 RAA
71  1985    1  1092 RAA
72  1986    1  1513 RAA

```

To transform the long table of the RAA data into a triangle we use the function `as.triangle`. The arguments we have to specify are the column names of the origin and development period and further the column which contains the values:

```
R> raa.tri <- as.triangle(raa,
                          origin="origin",
                          dev="dev",
                          value="value")
R> raa.tri
```

```

      dev
origin  1   2   3   4   5   6   7   8   9  10
1981 5012 3257 2638 898 1734 2642 1828 599 54 172
1982  106 4179 1111 5270 3116 1817 -103 673 535 NA
1983 3410 5582 4881 2268 2594 3479 649 603 NA NA

```

1984	5655	5900	4211	5500	2159	2658	984	NA	NA	NA
1985	1092	8473	6271	6333	3786	225	NA	NA	NA	NA
1986	1513	4932	5257	1233	2917	NA	NA	NA	NA	NA
1987	557	3463	6926	1368	NA	NA	NA	NA	NA	NA
1988	1351	5596	6165	NA	NA	NA	NA	NA	NA	NA
1989	3133	2262	NA	NA	NA	NA	NA	NA	NA	NA
1990	2063	NA	NA	NA	NA	NA	NA	NA	NA	NA

We note that the data has been stored as an incremental data set. As mentioned above, we could now use the function `incr2cum` to transform the triangle into a cumulative format.

We can transform a triangle back into a data frame structure:

```
R> raa.df <- as.data.frame(raa.tri, na.rm=TRUE)
R> head(raa.df)
```

	origin	dev	value
1981-1	1981	1	5012
1982-1	1982	1	106
1983-1	1983	1	3410
1984-1	1984	1	5655
1985-1	1985	1	1092
1986-1	1986	1	1513

This is particularly helpful when you would like to store your results back into a data base. Figure 4 gives you an idea of a potential data flow between R and data bases.

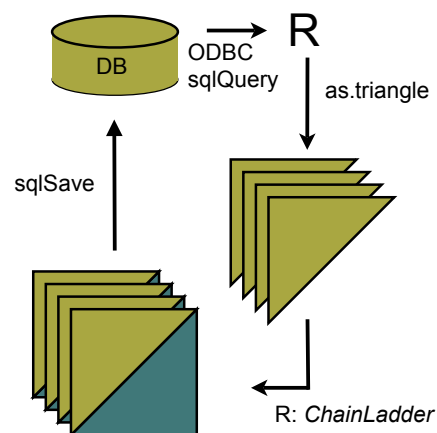


Figure 4: Flow chart of data between R and data bases.

4 Chain-ladder methods

The classical chain-ladder is a deterministic algorithm to forecast claims based on historical data. It assumes that the proportional developments of claims from one development period to the next are the same for all origin years.

4.1 Basic idea

Most commonly as a first step, the age-to-age link ratios are calculated as the volume weighted average development ratios of a cumulative loss development triangle from one development period to the next $C_{ik}, i, k = 1, \dots, n$.

$$f_k = \frac{\sum_{i=1}^{n-k} C_{i,k+1}}{\sum_{i=1}^{n-k} C_{i,k}} \quad (1)$$

```
R> n <- 10
R> f <- sapply(1:(n-1),
  function(i){
    sum(RAA[c(1:(n-i)),i+1])/sum(RAA[c(1:(n-i)),i])
  }
)
R> f

[1] 2.999 1.624 1.271 1.172 1.113 1.042 1.033 1.017 1.009
```

Often it is not suitable to assume that the oldest origin year is fully developed. A typical approach is to extrapolate the development ratios, e.g. assuming a log-linear model.

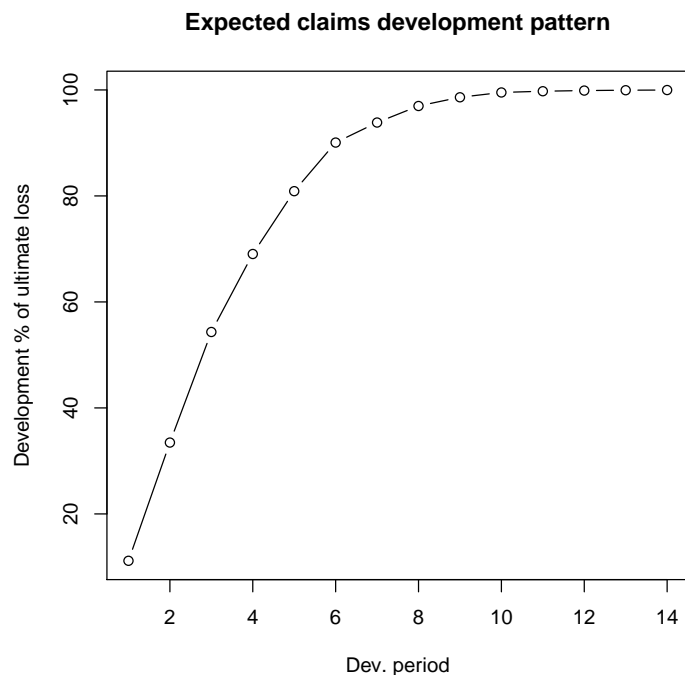
```
R> dev.period <- 1:(n-1)
R> plot(log(f-1) ~ dev.period, main="Log-linear extrapolation of age-to-age factors")
R> tail.model <- lm(log(f-1) ~ dev.period)
R> abline(tail.model)
R> co <- coef(tail.model)
R> ## extrapolate another 100 dev. period
R> tail <- exp(co[1] + c((n + 1):(n + 100)) * co[2]) + 1
R> f.tail <- prod(tail)
R> f.tail

[1] 1.005
```



The age-to-age factors allow us to plot the expected claims development patterns.

```
R> plot(100*(rev(1/cumprod(rev(c(f, tail[tail>1.0001]))))), t="b",
      main="Expected claims development pattern",
      xlab="Dev. period", ylab="Development % of ultimate loss")
```



The link ratios are then applied to the latest known cumulative claims amount to forecast the next development period. The *squaring* of the RAA triangle is calculated below, where an *ultimate* column is appended to the right to accommodate the expected development beyond the oldest age (10) of the triangle due to the tail factor (1.005) being greater than unity.

```
R> f <- c(f, f.tail)
R> fullRAA <- cbind(RAA, Ult = rep(0, 10))
R> for(k in 1:n){
  fullRAA[(n-k+1):n, k+1] <- fullRAA[(n-k+1):n, k]*f[k]
}
R> round(fullRAA)
```

	1	2	3	4	5	6	7	8	9	10	Ult
1981	5012	8269	10907	11805	13539	16181	18009	18608	18662	18834	18928
1982	106	4285	5396	10666	13782	15599	15496	16169	16704	16858	16942
1983	3410	8992	13873	16141	18735	22214	22863	23466	23863	24083	24204
1984	5655	11555	15766	21266	23425	26083	27067	27967	28441	28703	28847
1985	1092	9565	15836	22169	25955	26180	27278	28185	28663	28927	29072
1986	1513	6445	11702	12935	15852	17649	18389	19001	19323	19501	19599
1987	557	4020	10946	12314	14428	16064	16738	17294	17587	17749	17838
1988	1351	6947	13112	16664	19525	21738	22650	23403	23800	24019	24139


```
1989 3133 5395 8759 11132 13043 14521 15130 15634 15898 16045 16125
1990 2063 6188 10046 12767 14959 16655 17353 17931 18234 18402 18495
```

The total estimated outstanding loss under this method is about 53200:

```
R> sum(fullRAA[,11] - getLatestCumulative(RAA))
```

```
[1] 53202
```

This approach is also called Loss Development Factor (LDF) method.

More generally, the factors used to square the triangle need not always be drawn from the dollar weighted averages of the triangle. Other sources of factors from which the actuary may *select* link ratios include simple averages from the triangle, averages weighted toward more recent observations or adjusted for outliers, and benchmark patterns based on related, more credible loss experience. Also, since the ultimate value of claims is simply the product of the most current diagonal and the cumulative product of the link ratios, the completion of interior of the triangle is usually not displayed in favor of that multiplicative calculation.

For example, suppose the actuary decides that the volume weighted factors from the RAA triangle are representative of expected future growth, but discards the 1.005 tail factor derived from the loglinear fit in favor of a five percent tail (1.05) based on loss data from a larger book of similar business. The LDF method might be displayed in R as follows.

```
R> linkratios <- c(attr(ata(RAA), "vwtd"), tail = 1.05)
R> round(linkratios, 3) # display to only three decimal places
```

```
1-2 2-3 3-4 4-5 5-6 6-7 7-8 8-9 9-10 tail
2.999 1.624 1.271 1.172 1.113 1.042 1.033 1.017 1.009 1.050
```

```
R> LDF <- rev(cumprod(rev(linkratios)))
R> names(LDF) <- colnames(RAA) # so the display matches the triangle
R> round(LDF, 3)
```

```
1 2 3 4 5 6 7 8 9 10
9.366 3.123 1.923 1.513 1.292 1.160 1.113 1.078 1.060 1.050
```

```
R> currentEval <- getLatestCumulative(RAA)
R> # Reverse the LDFs so the first, least mature factor [1]
R> # is applied to the last origin year (1990)
R> EstdUlt <- currentEval * rev(LDF) #
R> # Start with the body of the exhibit
R> Exhibit <- data.frame(currentEval, LDF = round(rev(LDF), 3), EstdUlt)
```

```
R> # Tack on a Total row
R> Exhibit <- rbind(Exhibit,
  data.frame(currentEval=sum(currentEval), LDF=NA, EstdUlt=sum(EstdUlt),
    row.names = "Total"))
R> Exhibit
```

	currentEval	LDF	EstdUlt
1981	18834	1.050	19776
1982	16704	1.060	17701
1983	23466	1.078	25288
1984	27067	1.113	30138
1985	26180	1.160	30373
1986	15852	1.292	20476
1987	12314	1.513	18637
1988	13112	1.923	25220
1989	5395	3.123	16847
1990	2063	9.366	19323
Total	160987	NA	223778

Since the early 1990s several papers have been published to embed the simple chain-ladder method into a statistical framework. Ben Zehnwirth and Glenn Barnett point out in [ZB00] that the age-to-age link ratios can be regarded as the coefficients of a weighted linear regression through the origin, see also [Mur94].

```
R> lmCL <- function(i, Triangle){
  lm(y~x+0, weights=1/Triangle[,i],
    data=data.frame(x=Triangle[,i], y=Triangle[,i+1]))
}
R> sapply(lapply(c(1:(n-1)), lmCL, RAA), coef)
```

	x	x	x	x	x	x	x	x
2.999	1.624	1.271	1.172	1.113	1.042	1.033	1.017	1.009

4.2 Mack chain-ladder

Thomas Mack published in 1993 [Mac93] a method which estimates the standard errors of the chain-ladder forecast without assuming a distribution under three conditions.

Following the notation of Mack [Mac99] let C_{ik} denote the cumulative loss amounts of origin period (e.g. accident year) $i = 1, \dots, m$, with losses known for development period (e.g. development year) $k \leq n + 1 - i$.

In order to forecast the amounts C_{ik} for $k > n + 1 - i$ the Mack chain-ladder-model

assumes:

$$\text{CL1: } E[F_{ik}|C_{i1}, C_{i2}, \dots, C_{ik}] = f_k \text{ with } F_{ik} = \frac{C_{i,k+1}}{C_{ik}} \quad (2)$$

$$\text{CL2: } \text{Var}\left(\frac{C_{i,k+1}}{C_{ik}}|C_{i1}, C_{i2}, \dots, C_{ik}\right) = \frac{\sigma_k^2}{w_{ik}C_{ik}^\alpha} \quad (3)$$

$$\text{CL3: } \{C_{i1}, \dots, C_{in}\}, \{C_{j1}, \dots, C_{jn}\}, \text{ are independent for origin period } i \neq j \quad (4)$$

with $w_{ik} \in [0; 1], \alpha \in \{0, 1, 2\}$. If these assumptions hold, the Mack-chain-ladder-model gives an unbiased estimator for IBNR (Incurred But Not Reported) claims.

The Mack-chain-ladder model can be regarded as a weighted linear regression through the origin for each development period: $\text{lm}(y \sim x + 0, \text{weights}=w/x^{(2-\alpha)})$, where y is the vector of claims at development period $k+1$ and x is the vector of claims at development period k .

The Mack method is implemented in the ChainLadder package via the function `MackChainLadder`.

As an example we apply the `MackChainLadder` function to our triangle RAA:

```
R> mack <- MackChainLadder(RAA, est.sigma="Mack")
R> mack
```

```
MackChainLadder(Triangle = RAA, est.sigma = "Mack")
```

	Latest	Dev.To.Date	Ultimate	IBNR	Mack.S.E	CV(IBNR)
1981	18,834	1.000	18,834	0	0	NaN
1982	16,704	0.991	16,858	154	206	1.339
1983	23,466	0.974	24,083	617	623	1.010
1984	27,067	0.943	28,703	1,636	747	0.457
1985	26,180	0.905	28,927	2,747	1,469	0.535
1986	15,852	0.813	19,501	3,649	2,002	0.549
1987	12,314	0.694	17,749	5,435	2,209	0.406
1988	13,112	0.546	24,019	10,907	5,358	0.491
1989	5,395	0.336	16,045	10,650	6,333	0.595
1990	2,063	0.112	18,402	16,339	24,566	1.503

```
Totals
Latest: 160,987.00
Dev: 0.76
Ultimate: 213,122.23
IBNR: 52,135.23
Mack.S.E 26,909.01
CV(IBNR): 0.52
```

We can access the loss development factors and the full triangle via

```
R> mack$f
```

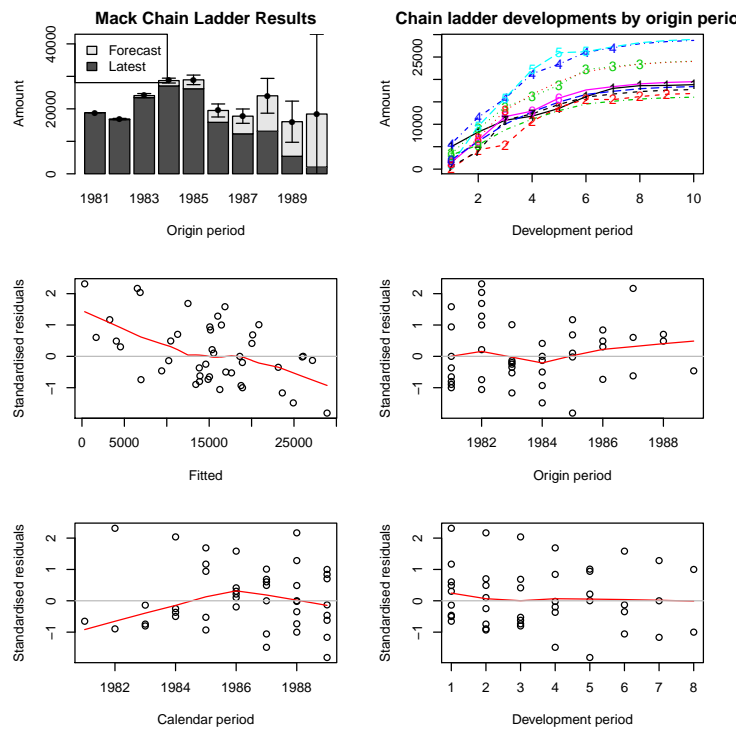
```
[1] 2.999 1.624 1.271 1.172 1.113 1.042 1.033 1.017 1.009 1.000
```

```
R> mack$FullTriangle
```

	dev									
origin	1	2	3	4	5	6	7	8	9	10
1981	5012	8269	10907	11805	13539	16181	18009	18608	18662	18834
1982	106	4285	5396	10666	13782	15599	15496	16169	16704	16858
1983	3410	8992	13873	16141	18735	22214	22863	23466	23863	24083
1984	5655	11555	15766	21266	23425	26083	27067	27967	28441	28703
1985	1092	9565	15836	22169	25955	26180	27278	28185	28663	28927
1986	1513	6445	11702	12935	15852	17649	18389	19001	19323	19501
1987	557	4020	10946	12314	14428	16064	16738	17294	17587	17749
1988	1351	6947	13112	16664	19525	21738	22650	23403	23800	24019
1989	3133	5395	8759	11132	13043	14521	15130	15634	15898	16045
1990	2063	6188	10046	12767	14959	16655	17353	17931	18234	18402

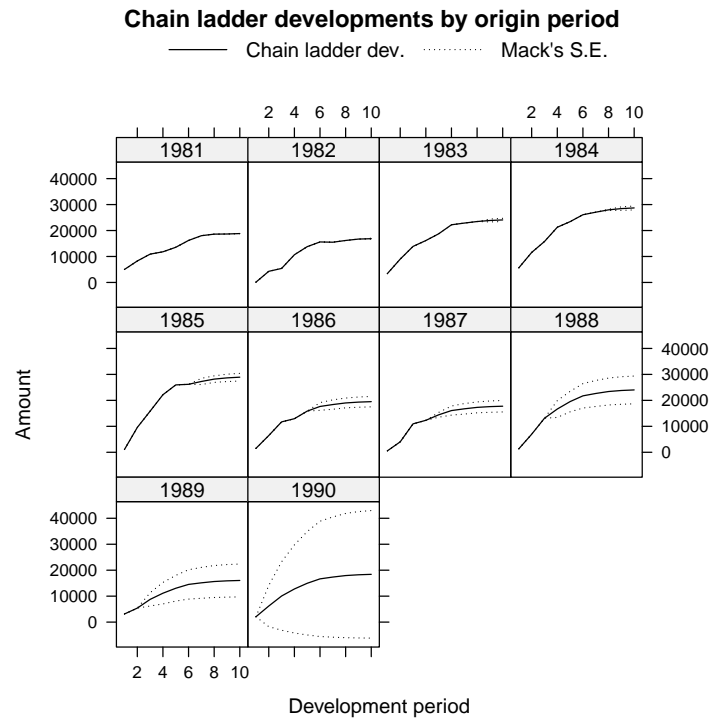
To check that Mack's assumption are valid review the residual plots, you should see no trends in either of them.

```
R> plot(mack)
```



We can plot the development, including the forecast and estimated standard errors by origin period by setting the argument `lattice=TRUE`.

```
R> plot(mack, lattice=TRUE)
```



4.3 Munich chain-ladder

Munich chain-ladder is a reserving method that reduces the gap between IBNR projections based on paid losses and IBNR projections based on incurred losses. The Munich chain-ladder method uses correlations between paid and incurred losses of the historical data into the projection for the future. [QM04].

```
R> MCLpaid
```

	dev						
origin	1	2	3	4	5	6	7
1	576	1804	1970	2024	2074	2102	2131
2	866	1948	2162	2232	2284	2348	NA
3	1412	3758	4252	4416	4494	NA	NA
4	2286	5292	5724	5850	NA	NA	NA
5	1868	3778	4648	NA	NA	NA	NA
6	1442	4010	NA	NA	NA	NA	NA
7	2044	NA	NA	NA	NA	NA	NA

```
R> MCLincurred
```

	dev						
origin	1	2	3	4	5	6	7
1	978	2104	2134	2144	2174	2182	2174
2	1844	2552	2466	2480	2508	2454	NA
3	2904	4354	4698	4600	4644	NA	NA
4	3502	5958	6070	6142	NA	NA	NA
5	2812	4882	4852	NA	NA	NA	NA
6	2642	4406	NA	NA	NA	NA	NA
7	5022	NA	NA	NA	NA	NA	NA

```
R> op <- par(mfrow=c(1,2))
R> plot(MCLpaid)
R> plot(MCLincurred)
R> par(op)
R> # Following the example in Quarg's (2004) paper:
R> MCL <- MunichChainLadder(MCLpaid, MCLincurred, est.sigmaP=0.1, est.sigmaI=0.1)
R> MCL
```

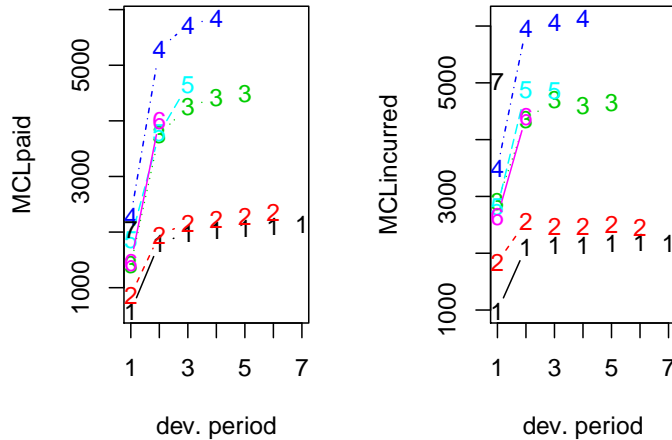
```
MunichChainLadder(Paid = MCLpaid, Incurred = MCLincurred, est.sigmaP = 0.1,
  est.sigmaI = 0.1)
```

	Latest Paid	Latest Incurred	Latest P/I Ratio	Ult. Paid	Ult. Incurred
1	2,131	2,174	0.980	2,131	2,174
2	2,348	2,454	0.957	2,383	2,444
3	4,494	4,644	0.968	4,597	4,629
4	5,850	6,142	0.952	6,119	6,176
5	4,648	4,852	0.958	4,937	4,950
6	4,010	4,406	0.910	4,656	4,665
7	2,044	5,022	0.407	7,549	7,650

	Ult. P/I Ratio
1	0.980
2	0.975
3	0.993
4	0.991
5	0.997
6	0.998
7	0.987

Totals	Paid	Incurred	P/I Ratio
Latest:	25,525	29,694	0.86
Ultimate:	32,371	32,688	0.99

```
R> plot(MCL)
```



4.4 Bootstrap chain-ladder

The `BootChainLadder` function uses a two-stage bootstrapping/simulation approach following the paper by England and Verrall [PR02]. In the first stage an ordinary chain-ladder methods is applied to the cumulative claims triangle. From this we calculate the scaled Pearson residuals which we bootstrap R times to forecast future incremental claims payments via the standard chain-ladder method. In the second stage we simulate the process error with the bootstrap value as the mean and using the process distribution assumed. The set of reserves obtained in this way forms the predictive distribution, from which summary statistics such as mean, prediction error or quantiles can be derived.

```
R> ## See also the example in section 8 of England & Verrall (2002)
R> ## on page 55.
R> B <- BootChainLadder(RAA, R=999, process.distr="gamma")
R> B
```

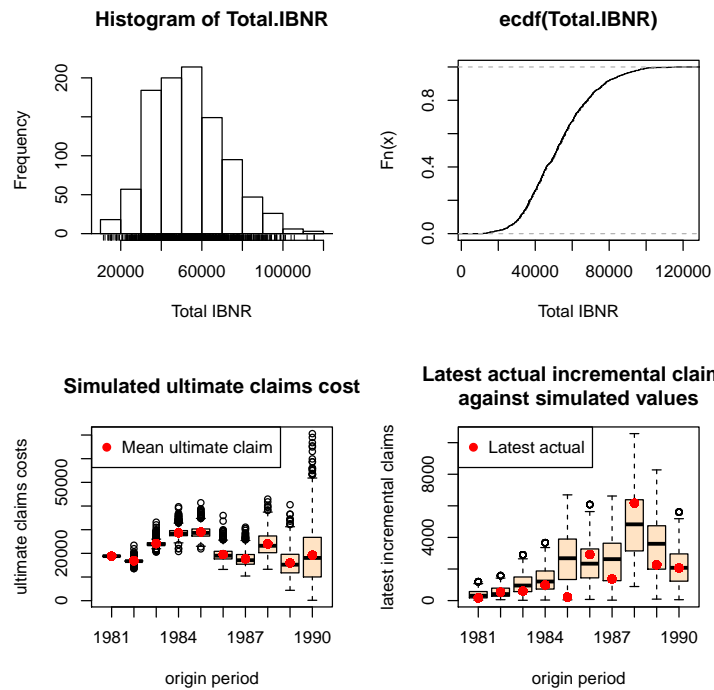
```
BootChainLadder(Triangle = RAA, R = 999, process.distr = "gamma")
```

	Latest	Mean	Ultimate	Mean	IBNR	IBNR.S.E	IBNR 75%	IBNR 95%
1981	18,834		18,834		0	0	0	0
1982	16,704		16,871		167	700	189	1,385
1983	23,466		24,116		650	1,262	1,025	3,077
1984	27,067		28,750		1,683	1,967	2,551	5,447
1985	26,180		29,041		2,861	2,424	4,127	7,487

1986	15,852	19,490	3,638	2,530	5,015	7,887
1987	12,314	17,720	5,406	3,145	7,194	11,344
1988	13,112	24,002	10,890	4,928	14,240	19,593
1989	5,395	15,982	10,587	5,763	14,200	21,041
1990	2,063	19,212	17,149	13,161	24,342	42,654

	Totals
Latest:	160,987
Mean Ultimate:	214,018
Mean IBNR:	53,031
IBNR.S.E	18,092
Total IBNR 75%:	64,072
Total IBNR 95%:	86,291

`R> plot(B)`



Quantiles of the bootstrap IBNR can be calculated via the `quantile` function:

`R> quantile(B, c(0.75,0.95,0.99, 0.995))`

`$ByOrigin`
 IBNR 75% IBNR 95% IBNR 99% IBNR 99.5%

1981	0.0	0	0	0
1982	188.8	1385	2756	3174
1983	1024.9	3077	5446	6048
1984	2551.2	5447	8362	9055
1985	4126.7	7487	10200	11179
1986	5015.3	7887	11973	13697
1987	7193.8	11344	14644	16256
1988	14240.3	19593	23658	25248
1989	14200.5	21041	25874	28811
1990	24342.0	42654	53955	61007

\$Totals

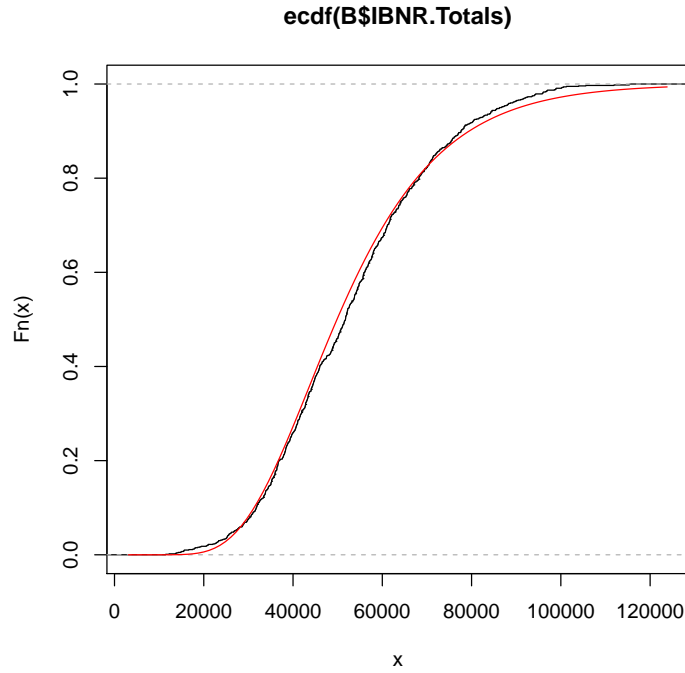
	Totals
IBNR 75%:	64072
IBNR 95%:	86291
IBNR 99%:	98869
IBNR 99.5%:	101269

The distribution of the IBNR appears to follow a log-normal distribution, so let's fit it:

```
R> ## fit a distribution to the IBNR
R> library(MASS)
R> plot(ecdf(B$IBNR.Totals))
R> ## fit a log-normal distribution
R> fit <- fitdistr(B$IBNR.Totals[B$IBNR.Totals>0], "lognormal")
R> fit
```

meanlog	sdlog
10.816666	0.363746
(0.011508)	(0.008138)

```
R> curve(plnorm(x,fit$estimate["meanlog"], fit$estimate["sdlog"]),
        col="red", add=TRUE)
```



4.5 Multivariate chain-ladder

The Mack chain ladder technique can be generalized to the multivariate setting where multiple reserving triangles are modelled and developed simultaneously. The advantage of the multivariate modelling is that correlations among different triangles can be modelled, which will lead to more accurate uncertainty assessments. Reserving methods that explicitly model the between-triangle contemporaneous correlations can be found in [PS05, MW08b]. Another benefit of multivariate loss reserving is that structural relationships between triangles can also be reflected, where the development of one triangle depends on past losses from other triangles. For example, there is generally need for the joint development of the paid and incurred losses [QM04]. Most of the chain-ladder-based multivariate reserving models can be summarised as sequential seemingly unrelated regressions [Zha10]. We note another strand of multivariate loss reserving builds a hierarchical structure into the model to allow estimation of one triangle to “borrow strength” from other triangles, reflecting the core insight of actuarial credibility [ZDG12].

Denote $Y_{i,k} = (Y_{i,k}^{(1)}, \dots, Y_{i,k}^{(N)})$ as an $N \times 1$ vector of cumulative losses at accident year i and development year k where (n) refers to the n -th triangle. [Zha10] specifies

the model in development period k as:

$$Y_{i,k+1} = A_k + B_k \cdot Y_{i,k} + \epsilon_{i,k}, \quad (5)$$

where A_k is a column of intercepts and B_k is the development matrix for development period k . Assumptions for this model are:

$$E(\epsilon_{i,k} | Y_{i,1}, \dots, Y_{i,I+1-k}) = 0. \quad (6)$$

$$\text{cov}(\epsilon_{i,k} | Y_{i,1}, \dots, Y_{i,I+1-k}) = D(Y_{i,k}^{-\delta/2}) \Sigma_k D(Y_{i,k}^{-\delta/2}). \quad (7)$$

$$\text{losses of different accident years are independent.} \quad (8)$$

$$\epsilon_{i,k} \text{ are symmetrically distributed.} \quad (9)$$

In the above, D is the diagonal operator, and δ is a known positive value that controls how the variance depends on the mean (as weights). This model is referred to as the general multivariate chain ladder [GMCL] in [Zha10]. A important special case where $A_k = 0$ and B_k 's are diagonal is a naive generalization of the chain ladder, often referred to as the multivariate chain ladder [MCL] [PS05].

In the following, we first introduce the class "triangles", for which we have defined several utility functions. Indeed, any input triangles to the `MultiChainLadder` function will be converted to "triangles" internally. We then present loss reserving methods based on the MCL and GMCL models in turn.

4.6 The "triangles" class

Consider the two liability loss triangles from [MW08b]. It comes as a list of two matrices :

```
R> str(liab)
```

```
List of 2
 $ GeneralLiab: num [1:14, 1:14] 59966 49685 51914 84937 98921 ...
 $ AutoLiab   : num [1:14, 1:14] 114423 152296 144325 145904 170333 ...
```

We can convert a list to a "triangles" object using

```
R> liab2 <- as(liab, "triangles")
R> class(liab2)
```

```
[1] "triangles"
attr(,"package")
[1] "ChainLadder"
```

We can find out what methods are available for this class:

```
R> showMethods(classes = "triangles")
```

For example, if we want to extract the last three columns of each triangle, we can use the "[" operator as follows:

```
R> # use drop = TRUE to remove rows that are all NA's
R> liab2[, 12:14, drop = TRUE]
```

An object of class "triangles"

```
[[1]]
      [,1] [,2] [,3]
[1,] 540873 547696 549589
[2,] 563571 562795    NA
[3,] 602710    NA    NA
```

```
[[2]]
      [,1] [,2] [,3]
[1,] 391328 391537 391428
[2,] 485138 483974    NA
[3,] 540742    NA    NA
```

The following combines two columns of the triangles to form a new matrix:

```
R> cbind2(liab2[1:3, 12])
```

```
      [,1] [,2]
[1,] 540873 391328
[2,] 563571 485138
[3,] 602710 540742
```

4.7 Separate chain ladder ignoring correlations

The form of regression models used in estimating the development parameters is controlled by the `fit.method` argument. If we specify `fit.method = "OLS"`, the ordinary least squares will be used and the estimation of development factors for each triangle is independent of the others. In this case, the residual covariance matrix Σ_k is diagonal. As a result, the multivariate model is equivalent to running multiple Mack chain ladders separately.

```
R> fit1 <- MultiChainLadder(liab, fit.method = "OLS")
R> lapply(summary(fit1)$report.summary, "[", 15, )
```

```
$`Summary Statistics for Triangle 1`
      Latest Dev.To.Date Ultimate IBNR S.E CV
```

```
Total 11343397      0.6482 17498658 6155261 427289 0.0694
```

```
$`Summary Statistics for Triangle 2`
```

```
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 8759806      0.8093 10823418 2063612 162872 0.0789
```

```
$`Summary Statistics for Triangle 1+2`
```

```
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 20103203      0.7098 28322077 8218874 457278 0.0556
```

In the above, we only show the total reserve estimate for each triangle to reduce the output. The full summary including the estimate for each year can be retrieved using the usual `summary` function. By default, the `summary` function produces reserve statistics for all individual triangles, as well as for the portfolio that is assumed to be the sum of the two triangles. This behaviour can be changed by supplying the `portfolio` argument. See the documentation for details.

We can verify if this is indeed the same as the univariate Mack chain ladder. For example, we can apply the `MackChainLadder` function to each triangle:

```
R> fit <- lapply(liab, MackChainLadder, est.sigma = "Mack")
R> # the same as the first triangle above
R> lapply(fit, function(x) t(summary(x)$Totals))
```

```
$GeneralLiab
```

```
      Latest:   Dev: Ultimate:   IBNR: Mack S.E.: CV(IBNR):
Totals 11343397 0.6482 17498658 6155261      427289 0.06942
```

```
$AutoLiab
```

```
      Latest:   Dev: Ultimate:   IBNR: Mack S.E.: CV(IBNR):
Totals 8759806 0.8093 10823418 2063612      162872 0.07893
```

The argument `mse.method` controls how the mean square errors are computed. By default, it implements the Mack method. An alternative method is the conditional re-sampling approach in [BBMW06], which assumes the estimated parameters are independent. This is used when `mse.method = "Independence"`. For example, the following reproduces the result in [BBMW06]. Note that the first argument must be a list, even though only one triangle is used.

```
R> (B1 <- MultiChainLadder(list(GenIns), fit.method = "OLS",
  mse.method = "Independence"))
```

```
$`Summary Statistics for Input Triangle`
```

```
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
1      3,901,463      1.0000 3,901,463      0      0 0.000
```

2	5,339,085	0.9826	5,433,719	94,634	75,535	0.798
3	4,909,315	0.9127	5,378,826	469,511	121,700	0.259
4	4,588,268	0.8661	5,297,906	709,638	133,551	0.188
5	3,873,311	0.7973	4,858,200	984,889	261,412	0.265
6	3,691,712	0.7223	5,111,171	1,419,459	411,028	0.290
7	3,483,130	0.6153	5,660,771	2,177,641	558,356	0.256
8	2,864,498	0.4222	6,784,799	3,920,301	875,430	0.223
9	1,363,294	0.2416	5,642,266	4,278,972	971,385	0.227
10	344,014	0.0692	4,969,825	4,625,811	1,363,385	0.295
Total	34,358,090	0.6478	53,038,946	18,680,856	2,447,618	0.131

4.8 Multivariate chain ladder using seemingly unrelated regressions

To allow correlations to be incorporated, we employ the seemingly unrelated regressions (see the package `systemfit`) that simultaneously model the two triangles in each development period. This is invoked when we specify `fit.method = "SUR"`:

```
R> fit2 <- MultiChainLadder(liab, fit.method = "SUR")
R> lapply(summary(fit2)$report.summary, "[", 15, )
```

```
$`Summary Statistics for Triangle 1`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 11343397      0.6484 17494907 6151510 419293 0.0682
```

```
$`Summary Statistics for Triangle 2`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total  8759806      0.8095 10821341 2061535 162464 0.0788
```

```
$`Summary Statistics for Triangle 1+2`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 20103203      0.71 28316248 8213045 500607 0.061
```

We see that the portfolio prediction error is inflated to 500,607 from 457,278 in the separate development model ("OLS"). This is because of the positive correlation between the two triangles. The estimated correlation for each development period can be retrieved through the `residCor` function:

```
R> round(unlist(residCor(fit2)), 3)

[1] 0.247 0.495 0.682 0.446 0.487 0.451 -0.172 0.805 0.337 0.688
[11] -0.004 1.000 0.021
```

Similarly, most methods that work for linear models such as `coef`, `fitted`, `resid` and so on will also work. Since we have a sequence of models, the retrieved results

from these methods are stored in a list. For example, we can retrieve the estimated development factors for each period as

```
R> do.call("rbind", coef(fit2))
```

	eq1_x[[1]]	eq2_x[[2]]
[1,]	3.227	2.2224
[2,]	1.719	1.2688
[3,]	1.352	1.1200
[4,]	1.179	1.0665
[5,]	1.106	1.0356
[6,]	1.055	1.0168
[7,]	1.026	1.0097
[8,]	1.015	1.0002
[9,]	1.012	1.0038
[10,]	1.006	0.9994
[11,]	1.005	1.0039
[12,]	1.005	0.9989
[13,]	1.003	0.9997

The smaller-than-one development factors after the 10-th period for the second triangle indeed result in negative IBNR estimates for the first several accident years in that triangle.

The package also offers the `plot` method that produces various summary and diagnostic figures:

```
R> parold <- par(mfrow = c(4, 2), mar = c(4, 4, 2, 1),
  mgp = c(1.3, 0.3, 0), tck = -0.02)
R> plot(fit2, which.triangle = 1:2, which.plot = 1:4)
R> par(parold)
```

The resulting plots are shown in Figure 5. We use `which.triangle` to suppress the plot for the portfolio, and use `which.plot` to select the desired types of plots. See the documentation for possible values of these two arguments.

4.9 Other residual covariance estimation methods

Internally, the `MultiChainLadder` calls the `systemfit` function to fit the regression models period by period. When SUR models are specified, there are several ways to estimate the residual covariance matrix Σ_k . Available methods are "noDfCor", "geomean", "max", and "Theil" with the default as "geomean". The method "Theil" will produce unbiased covariance estimate, but the resulting estimate may not be positive semi-definite. This is also the estimator used by [MW08b]. However, this method does not work out of the box for the `liab` data, and is perhaps one

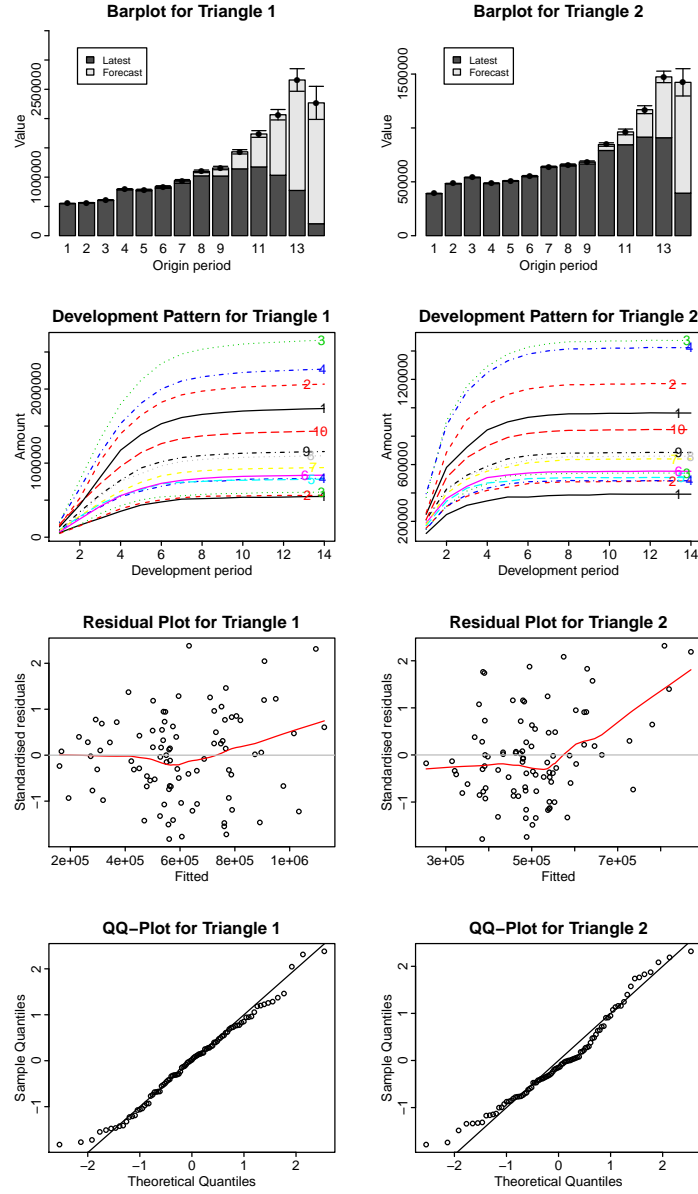


Figure 5: Summary and diagnostic plots from a MultiChainLadder object.

of the reasons [MW08b] used extrapolation to get the estimate for the last several periods.

Indeed, for most applications, we recommend the use of separate chain ladders for the tail periods to stabilize the estimation - there are few data points in the tail and running a multivariate model often produces extremely volatile estimates or even fails. To facilitate such an approach, the package offers the `MultiChainLadder2` function, which implements a split-and-join procedure: we split the input data into two parts, specify a multivariate model with rich structures on the first part (with enough data) to reflect the multivariate dependencies, apply separate univariate chain ladders on the second part, and then join the two models together to produce the final predictions. The splitting is determined by the "last" argument, which specifies how many of the development periods in the tail go into the second part of the split. The type of the model structure to be specified for the first part of the split model in `MultiChainLadder2` is controlled by the `type` argument. It takes one of the following values: "MCL"- the multivariate chain ladder with diagonal development matrix; "MCL+int"- the multivariate chain ladder with additional intercepts; "GMCL-int"- the general multivariate chain ladder without intercepts; and "GMCL" - the full general multivariate chain ladder with intercepts and non-diagonal development matrix.

For example, the following fits the SUR method to the first part (the first 11 columns) using the unbiased residual covariance estimator in [MW08b], and separate chain ladders for the rest:

```
R> require(systemfit)
R> W1 <- MultiChainLadder2(liab, mse.method = "Independence",
                           control = systemfit.control(methodResidCov = "Theil"))
R> lapply(summary(W1)$report.summary, "[", 15, )
```

```
$`Summary Statistics for Triangle 1`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 11343397      0.6483 17497403 6154006 427041 0.0694
```

```
$`Summary Statistics for Triangle 2`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total  8759806      0.8095 10821034 2061228 162785 0.079
```

```
$`Summary Statistics for Triangle 1+2`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 20103203      0.7099 28318437 8215234 505376 0.0615
```

Similarly, the iterative residual covariance estimator in [MW08b] can also be used, in which we use the control parameter `maxiter` to determine the number of iterations:

```
R> for (i in 1:5){
  W2 <- MultiChainLadder2(liab, mse.method = "Independence",
```

```

        control = systemfit.control(methodResidCov = "Theil", maxiter = i))
    print(format(summary(W2)$report.summary[[3]][15, 4:5],
              digits = 6, big.mark = ","))
}

```

```

          IBNR      S.E
Total 8,215,234 505,376
          IBNR      S.E
Total 8,215,357 505,443
          IBNR      S.E
Total 8,215,362 505,444
          IBNR      S.E
Total 8,215,362 505,444
          IBNR      S.E
Total 8,215,362 505,444

```

```
R> lapply(summary(W2)$report.summary, "[", 15, )
```

```

$`Summary Statistics for Triangle 1`
      Latest Dev.To.Date Ultimate      IBNR      S.E      CV
Total 11343397      0.6483 17497526 6154129 427074 0.0694

```

```

$`Summary Statistics for Triangle 2`
      Latest Dev.To.Date Ultimate      IBNR      S.E      CV
Total 8759806      0.8095 10821039 2061233 162790 0.079

```

```

$`Summary Statistics for Triangle 1+2`
      Latest Dev.To.Date Ultimate      IBNR      S.E      CV
Total 20103203      0.7099 28318565 8215362 505444 0.0615

```

We see that the covariance estimate converges in three steps. These are very similar to the results in [MW08b], the small difference being a result of the different approaches used in the last three periods.

Also note that in the above two examples, the argument `control` is not defined in the prototype of the `MultiChainLadder`. It is an argument that is passed to the `systemfit` function through the `...` mechanism. Users are encouraged to explore how other options available in `systemfit` can be applied.

4.10 Model with intercepts

Consider the auto triangles from [Zha10]. It includes three automobile insurance triangles: personal auto paid, personal auto incurred, and commercial auto paid.

```
R> str(auto)
```

List of 3

```
$ PersonalAutoPaid      : num [1:10, 1:10] 101125 102541 114932 114452 115597 ...
$ PersonalAutoIncurred: num [1:10, 1:10] 325423 323627 358410 405319 434065 ...
$ CommercialAutoPaid   : num [1:10, 1:10] 19827 22331 22533 23128 25053 ...
```

It is a reasonable expectation that these triangles will be correlated. So we run a MCL model on them:

```
R> f0 <- MultiChainLadder2(auto, type = "MCL")
R> # show correlation- the last three columns have zero correlation
R> # because separate chain ladders are used
R> print(do.call(cbind, residCor(f0)), digits = 3)
```

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]
(1,2)	0.327	-0.0101	0.598	0.711	0.8565	0.928	0	0	0
(1,3)	0.870	0.9064	0.939	0.261	-0.0607	0.911	0	0	0
(2,3)	0.198	-0.3217	0.558	0.380	0.3586	0.931	0	0	0

However, from the residual plot, the first row in Figure 6, it is evident that the default mean structure in the MCL model is not adequate. Usually this is a common problem with the chain ladder based models, owing to the missing of intercepts.

We can improve the above model by including intercepts in the SUR fit as follows:

```
R> f1 <- MultiChainLadder2(auto, type = "MCL+int")
```

The corresponding residual plot is shown in the second row in Figure 6. We see that these residuals are randomly scattered around zero and there is no clear pattern compared to the plot from the MCL model.

The default summary computes the portfolio estimates as the sum of all the triangles. This is not desirable because the first two triangles are both from the personal auto line. We can overwrite this via the `portfolio` argument. For example, the following uses the two paid triangles as the portfolio estimate:

```
R> lapply(summary(f1, portfolio = "1+3")@report.summary, "[", 11, )
```

```
$`Summary Statistics for Triangle 1`
      Latest Dev.To.Date Ultimate  IBNR   S.E    CV
Total 3290539      0.8537  3854572 564033 19089 0.0338
```

```
$`Summary Statistics for Triangle 2`
      Latest Dev.To.Date Ultimate  IBNR   S.E    CV
Total 3710614      0.9884  3754197 43583 18839 0.4323
```

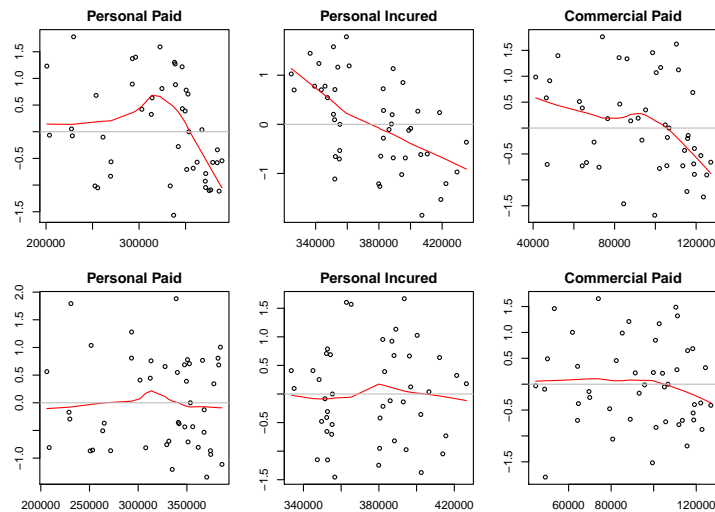


Figure 6: Residual plots for the MCL model (first row) and the GMCL (MCL+int) model (second row) for the auto data.

```
$`Summary Statistics for Triangle 3`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 1043851      0.7504  1391064 347213 27716 0.0798
```

```
$`Summary Statistics for Triangle 1+3`
      Latest Dev.To.Date Ultimate   IBNR   S.E   CV
Total 4334390      0.8263  5245636 911246 38753 0.0425
```

4.11 Joint modelling of the paid and incurred losses

Although the model with intercepts proved to be an improvement over the MCL model, it still fails to account for the structural relationship between triangles. In particular, it produces divergent paid-to-incurred loss ratios for the personal auto line:

```
R> ult <- summary(f1)$Ultimate
R> print(ult[, 1] /ult[, 2], 3)
```

	1	2	3	4	5	6	7	8	9	10	Total
	0.995	0.995	0.993	0.992	0.995	0.996	1.021	1.067	1.112	1.114	1.027

We see that for accident years 9-10, the paid-to-incurred loss ratios are more than 110%. This can be fixed by allowing the development of the paid/incurred triangles

to depend on each other. That is, we include the past values from the paid triangle as predictors when developing the incurred triangle, and vice versa.

We illustrate this ignoring the commercial auto triangle. See the demo for a model that uses all three triangles. We also include the MCL model and the Munich chain ladder as a comparison:

```
R> da <- auto[1:2]
R> # MCL with diagonal development
R> M0 <- MultiChainLadder(da)
R> # non-diagonal development matrix with no intercepts
R> M1 <- MultiChainLadder2(da, type = "GMCL-int")
R> # Munich Chain Ladder
R> M2 <- MunichChainLadder(da[[1]], da[[2]])
R> # compile results and compare projected paid to incurred ratios
R> r1 <- lapply(list(M0, M1), function(x){
  ult <- summary(x)@Ultimate
  ult[, 1] / ult[, 2]
})
R> names(r1) <- c("MCL", "GMCL")
R> r2 <- summary(M2)[[1]][, 6]
R> r2 <- c(r2, summary(M2)[[2]][2, 3])
R> print(do.call(cbind, c(r1, list(MuCl = r2))) * 100, digits = 4)
```

	MCL	GMCL	MuCl
1	99.50	99.50	99.50
2	99.49	99.49	99.55
3	99.29	99.29	100.23
4	99.20	99.20	100.23
5	99.83	99.56	100.04
6	100.43	99.66	100.03
7	103.53	99.76	99.95
8	111.24	100.02	99.81
9	122.11	100.20	99.67
10	126.28	100.18	99.69
Total	105.58	99.68	99.88

5 Clark's methods

The ChainLadder package contains functionality to carry out the methods described in the paper ⁶ by David Clark [Cla03]. Using a longitudinal analysis approach, Clark assumes that losses develop according to a theoretical *growth curve*. The LDF method is a special case of this approach where the growth curve can

⁶ This paper is on the CAS Exam 6 syllabus.

be considered to be either a step function or piecewise linear. Clark envisions a growth curve as measuring the percent of ultimate loss that can be expected to have emerged as of each age of an origin period. The paper describes two methods that fit this model.

The LDF method assumes that the ultimate losses in each origin period are separate and unrelated. The goal of the method, therefore, is to estimate parameters for the ultimate losses and for the growth curve in order to maximize the likelihood of having observed the data in the triangle.

The CapeCod method assumes that the *a priori* expected ultimate losses in each origin year are the product of earned premium that year and a theoretical loss ratio. The CapeCod method, therefore, need estimate potentially far fewer parameters: for the growth function and for the theoretical loss ratio.

One of the side benefits of using maximum likelihood to estimate parameters is that its associated asymptotic theory provides uncertainty estimates for the parameters. Observing that the reserve estimates by origin year are functions of the estimated parameters, uncertainty estimates of these functional values are calculated according to the *Delta method*, which is essentially a linearisation of the problem based on a Taylor series expansion.

The two functional forms for growth curves considered in Clark's paper are the log-logistic function (a.k.a., the inverse power curve) and the Weibull function, both being two-parameter functions. Clark uses the parameters ω and θ in his paper. Clark's methods work on incremental losses. His likelihood function is based on the assumption that incremental losses follow an over-dispersed Poisson (ODP) process.

5.1 Clark's LDF method

Consider again the RAA triangle. Accepting all defaults, the Clark LDF Method would estimate total ultimate losses of 272,009 and a reserve (FutureValue) of 111,022, or almost twice the value based on the volume weighted average link ratios and loglinear fit in section 3.2.1 above.

```
R> ClarkLDF(RAA)
```

Origin	CurrentValue	Ldf	UltimateValue	FutureValue	StdError	CV%
1981	18,834	1.216	22,906	4,072	2,792	68.6
1982	16,704	1.251	20,899	4,195	2,833	67.5
1983	23,466	1.297	30,441	6,975	4,050	58.1
1984	27,067	1.360	36,823	9,756	5,147	52.8
1985	26,180	1.451	37,996	11,816	5,858	49.6
1986	15,852	1.591	25,226	9,374	4,877	52.0
1987	12,314	1.829	22,528	10,214	5,206	51.0
1988	13,112	2.305	30,221	17,109	7,568	44.2
1989	5,395	3.596	19,399	14,004	7,506	53.6

1990	2,063	12.394	25,569	23,506	17,227	73.3
Total	160,987		272,009	111,022	36,102	32.5

Most of the difference is due to the heavy tail, 21.6%, implied by the inverse power curve fit. Clark recognizes that the log-logistic curve can take an unreasonably long length of time to flatten out. If according to the actuary's experience most claims close as of, say, 20 years, the growth curve can be truncated accordingly by using the maxage argument:

```
R> ClarkLDF(RAA, maxage = 20)
```

Origin	CurrentValue	Ldf	UltimateValue	FutureValue	StdError	CV%
1981	18,834	1.124	21,168	2,334	1,765	75.6
1982	16,704	1.156	19,314	2,610	1,893	72.6
1983	23,466	1.199	28,132	4,666	2,729	58.5
1984	27,067	1.257	34,029	6,962	3,559	51.1
1985	26,180	1.341	35,113	8,933	4,218	47.2
1986	15,852	1.471	23,312	7,460	3,775	50.6
1987	12,314	1.691	20,819	8,505	4,218	49.6
1988	13,112	2.130	27,928	14,816	6,300	42.5
1989	5,395	3.323	17,927	12,532	6,658	53.1
1990	2,063	11.454	23,629	21,566	15,899	73.7
Total	160,987		251,369	90,382	26,375	29.2

The Weibull growth curve tends to be faster developing than the log-logistic:

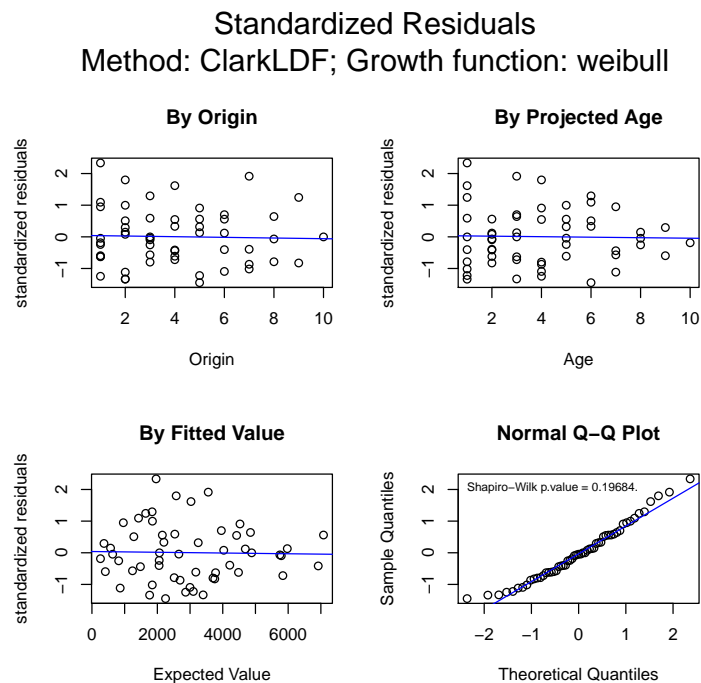
```
R> ClarkLDF(RAA, G="weibull")
```

Origin	CurrentValue	Ldf	UltimateValue	FutureValue	StdError	CV%
1981	18,834	1.022	19,254	420	700	166.5
1982	16,704	1.037	17,317	613	855	139.5
1983	23,466	1.060	24,875	1,409	1,401	99.4
1984	27,067	1.098	29,728	2,661	2,037	76.5
1985	26,180	1.162	30,419	4,239	2,639	62.2
1986	15,852	1.271	20,151	4,299	2,549	59.3
1987	12,314	1.471	18,114	5,800	3,060	52.8
1988	13,112	1.883	24,692	11,580	4,867	42.0
1989	5,395	2.988	16,122	10,727	5,544	51.7
1990	2,063	9.815	20,248	18,185	12,929	71.1
Total	160,987		220,920	59,933	19,149	32.0

It is recommend to inspect the residuals to help assess the reasonableness of the model relative to the actual data.

Although there is some evidence of heteroscedasticity with increasing ages and fitted values, the residuals otherwise appear randomly scattered around a horizontal line


```
R> plot(ClarkLDF(RAA, G="weibull"))
```



through the origin. The q-q plot shows evidence of a lack of fit in the tails, but the p-value of almost 0.2 can be considered too high to reject outright the assumption of normally distributed standardized residuals⁷.

5.2 Clark's Cap Cod method

The RAA data set, widely researched in the literature, has no premium associated with it traditionally. Let's assume a constant earned premium of 40000 each year, and a Weibull growth function:

```
R> ClarkCapeCod(RAA, Premium = 40000, G = "weibull")
```

Origin	CurrentValue	Premium	ELR	FutureGrowthFactor	FutureValue	UltimateValue
1981	18,834	40,000	0.566	0.0192	436	19,270
1982	16,704	40,000	0.566	0.0320	725	17,429
1983	23,466	40,000	0.566	0.0525	1,189	24,655

⁷As an exercise, the reader can confirm that the normal distribution assumption is rejected at the 5% level with the log-logistic curve.

1984	27,067	40,000	0.566	0.0848	1,921	28,988
1985	26,180	40,000	0.566	0.1345	3,047	29,227
1986	15,852	40,000	0.566	0.2093	4,741	20,593
1987	12,314	40,000	0.566	0.3181	7,206	19,520
1988	13,112	40,000	0.566	0.4702	10,651	23,763
1989	5,395	40,000	0.566	0.6699	15,176	20,571
1990	2,063	40,000	0.566	0.9025	20,444	22,507
Total	160,987	400,000			65,536	226,523
StdError	CV%					
692	158.6					
912	125.7					
1,188	99.9					
1,523	79.3					
1,917	62.9					
2,360	49.8					
2,845	39.5					
3,366	31.6					
3,924	25.9					
4,491	22.0					
12,713	19.4					

The estimated expected loss ratio is 0.566. The total outstanding loss is about 10% higher than with the LDF method. The standard error, however, is lower, probably due to the fact that there are fewer parameters to estimate with the CapeCod method, resulting in less parameter risk.

A plot of this model shows similar residuals By Origin and Projected Age to those from the LDF method, a better spread By Fitted Value, and a slightly better q-q plot, particularly in the upper tail.

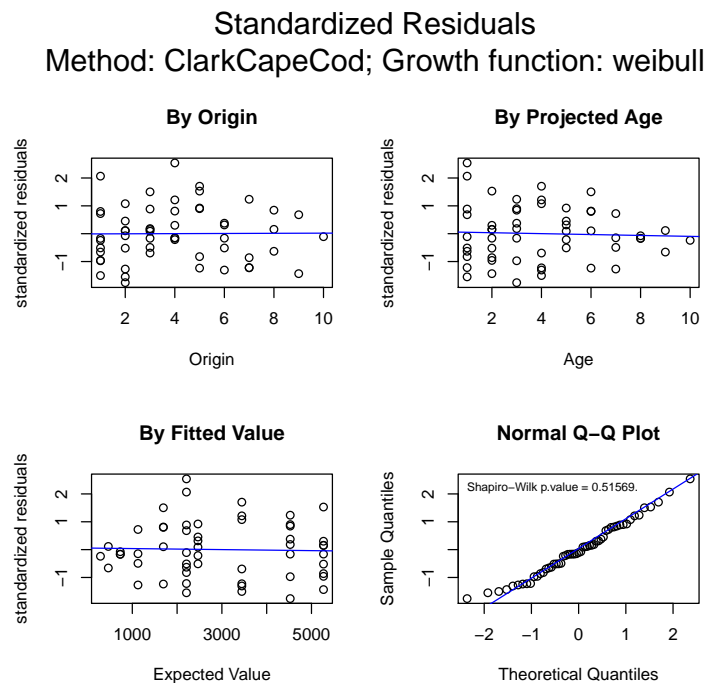
6 Generalised linear model methods

Recent years have also seen growing interest in using generalised linear models [GLM] for insurance loss reserving. The use of GLM in insurance loss reserving has many compelling aspects, e.g.,

- when over-dispersed Poisson model is used, it reproduces the estimates from Chain Ladder;
- it provides a more coherent modelling framework than the Mack method;
- all the relevant established statistical theory can be directly applied to perform hypothesis testing and diagnostic checking;

The `glmReserve` function takes an insurance loss triangle, converts it to incremental losses internally if necessary, transforms it to the long format (see `as.data.frame`)

```
R> plot(ClarkCapeCod(RAA, Premium = 40000, G = "weibull"))
```



and fits the resulting loss data with a generalised linear model where the mean structure includes both the accident year and the development lag effects. The function also provides both analytical and bootstrapping methods to compute the associated prediction errors. The bootstrapping approach also simulates the full predictive distribution, based on which the user can compute other uncertainty measures such as predictive intervals.

Only the Tweedie family of distributions are allowed, that is, the exponential family that admits a power variance function $V(\mu) = \mu^p$. The variance power p is specified in the `var.power` argument, and controls the type of the distribution. When the Tweedie compound Poisson distribution $1 < p < 2$ is to be used, the user has the option to specify `var.power = NULL`, where the variance power p will be estimated from the data using the `cplm` package [Zha12].

For example, the following fits the over-dispersed Poisson model and spells out the estimated reserve information:

```
R> # load data
R> data(GenIns)
R> GenIns <- GenIns / 1000
```

```
R> # fit Poisson GLM
R> (fit1 <- glmReserve(GenIns))
```

	Latest	Dev.To.Date	Ultimate	IBNR	S.E	CV
2	5339	0.98252	5434	95	110.1	1.1589
3	4909	0.91263	5379	470	216.0	0.4597
4	4588	0.86599	5298	710	260.9	0.3674
5	3873	0.79725	4858	985	303.6	0.3082
6	3692	0.72235	5111	1419	375.0	0.2643
7	3483	0.61527	5661	2178	495.4	0.2274
8	2864	0.42221	6784	3920	790.0	0.2015
9	1363	0.24162	5642	4279	1046.5	0.2446
10	344	0.06922	4970	4626	1980.1	0.4280
total	30457	0.61982	49138	18681	2945.7	0.1577

We can also extract the underlying GLM model by specifying type = "model" in the summary function:

```
R> summary(fit1, type = "model")
```

Call:

```
glm(formula = value ~ factor(origin) + factor(dev), family = fam,
     data = ldaFit, offset = offset)
```

Deviance Residuals:

Min	1Q	Median	3Q	Max
-14.701	-3.913	-0.688	3.675	15.633

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	5.59865	0.17292	32.38	< 2e-16 ***
factor(origin)2	0.33127	0.15354	2.16	0.0377 *
factor(origin)3	0.32112	0.15772	2.04	0.0492 *
factor(origin)4	0.30596	0.16074	1.90	0.0650 .
factor(origin)5	0.21932	0.16797	1.31	0.1999
factor(origin)6	0.27008	0.17076	1.58	0.1225
factor(origin)7	0.37221	0.17445	2.13	0.0398 *
factor(origin)8	0.55333	0.18653	2.97	0.0053 **
factor(origin)9	0.36893	0.23918	1.54	0.1317
factor(origin)10	0.24203	0.42756	0.57	0.5749
factor(dev)2	0.91253	0.14885	6.13	4.7e-07 ***
factor(dev)3	0.95883	0.15257	6.28	2.9e-07 ***
factor(dev)4	1.02600	0.15688	6.54	1.3e-07 ***
factor(dev)5	0.43528	0.18391	2.37	0.0234 *
factor(dev)6	0.08006	0.21477	0.37	0.7115

```

factor(dev)7      -0.00638    0.23829   -0.03    0.9788
factor(dev)8      -0.39445    0.31029   -1.27    0.2118
factor(dev)9       0.00938    0.32025    0.03    0.9768
factor(dev)10     -1.37991    0.89669   -1.54    0.1326
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

(Dispersion parameter for Tweedie family taken to be 52.6)

```

Null deviance: 10699 on 54 degrees of freedom
Residual deviance: 1903 on 36 degrees of freedom
AIC: NA

```

Number of Fisher Scoring iterations: 4

Similarly, we can fit the Gamma and a compound Poisson GLM reserving model by changing the `var.power` argument:

```

R> # Gamma GLM
R> (fit2 <- glmReserve(GenIns, var.power = 2))

```

	Latest	Dev.To.Date	Ultimate	IBNR	S.E	CV
2	5339	0.98288	5432	93	45.17	0.4857
3	4909	0.91655	5356	447	160.56	0.3592
4	4588	0.88248	5199	611	177.62	0.2907
5	3873	0.79611	4865	992	254.47	0.2565
6	3692	0.71757	5145	1453	351.33	0.2418
7	3483	0.61440	5669	2186	526.29	0.2408
8	2864	0.43870	6529	3665	941.32	0.2568
9	1363	0.24854	5485	4122	1175.95	0.2853
10	344	0.07078	4860	4516	1667.39	0.3692
total	30457	0.62742	48543	18086	2702.71	0.1494

```

R> # compound Poisson GLM (variance function estimated from the data):
R> #(fit3 <- glmReserve(GenIns, var.power = NULL))

```

By default, the formulaic approach is used to compute the prediction errors. We can also carry out bootstrapping simulations by specifying `mse.method = "bootstrap"` (note that this argument supports partial match):

```

R> set.seed(11)
R> (fit5 <- glmReserve(GenIns, mse.method = "boot"))

```

	Latest	Dev.To.Date	Ultimate	IBNR	S.E	CV
2	5339	0.98252	5434	95	105.4	1.1098

3	4909	0.91263	5379	470	216.1	0.4597
4	4588	0.86599	5298	710	266.6	0.3755
5	3873	0.79725	4858	985	307.5	0.3122
6	3692	0.72235	5111	1419	376.3	0.2652
7	3483	0.61527	5661	2178	496.1	0.2278
8	2864	0.42221	6784	3920	812.9	0.2074
9	1363	0.24162	5642	4279	1050.9	0.2456
10	344	0.06922	4970	4626	2004.1	0.4332
total	30457	0.61982	49138	18681	2959.4	0.1584

When bootstrapping is used, the resulting object has three additional components - "sims.par", "sims.reserve.mean", and "sims.reserve.pred" that store the simulated parameters, mean values and predicted values of the reserves for each year, respectively.

```
R> names(fit5)
```

```
[1] "call"           "summary"        "Triangle"
[4] "FullTriangle"   "model"          "sims.par"
[7] "sims.reserve.mean" "sims.reserve.pred"
```

We can thus compute the quantiles of the predictions based on the simulated samples in the "sims.reserve.pred" element as:

```
R> pr <- as.data.frame(fit5$sims.reserve.pred)
R> qv <- c(0.025, 0.25, 0.5, 0.75, 0.975)
R> res.q <- t(apply(pr, 2, quantile, qv))
R> print(format(round(res.q), big.mark = ","), quote = FALSE)
```

	2.5%	25%	50%	75%	97.5%
2	0	34	82	170	376
3	136	337	470	615	987
4	279	556	719	917	1,302
5	506	797	972	1,197	1,674
6	774	1,159	1,404	1,666	2,203
7	1,329	1,877	2,210	2,547	3,303
8	2,523	3,463	3,991	4,572	5,713
9	2,364	3,593	4,310	5,013	6,531
10	913	3,354	4,487	5,774	9,165

The full predictive distribution of the simulated reserves for each year can be visualized easily:

```
R> library(ggplot2)
R> library(reshape2)
```

```

R> prm <- melt(pr)
R> names(prm) <- c("year", "reserve")
R> gg <- ggplot(prm, aes(reserve))
R> gg <- gg + geom_density(aes(fill = year), alpha = 0.3) +
  facet_wrap(~year, nrow = 2, scales = "free") +
  theme(legend.position = "none")
R> print(gg)

```

7 Paid-incurred chain model

The Paid-incurred chain model was published by Merz and Wüthrich in 2010 [MW10]. It combines claims payments and incurred losses information in a mathematically rigorous and consistent way to get a unified ultimate loss prediction.

7.1 Model assumptions

The model assumptions for the Log-Normal PIC Model are the following:

- Conditionally, given $\Theta = (\Phi_0, \dots, \Phi_I, \Psi_0, \dots, \Psi_{I-1}, \sigma_0, \dots, \sigma_{I-1}, \tau_0, \dots, \tau_{I-1})$ we have
 - the random vector $(\xi_{0,0}, \dots, \xi_{I,I}, \zeta_{0,0}, \dots, \zeta_{I,I-1})$ has multivariate Gaussian distribution with uncorrelated components given by

$$\xi_{i,j} \sim N(\Phi_j, \sigma_j^2),$$

$$\zeta_{k,l} \sim N(\Psi_l, \tau_l^2);$$

- cumulative payments are given by the recursion

$$P_{i,j} = P_{i,j-1} \exp(\xi_{i,j}),$$

with initial value $P_{i,0} = \exp(\xi_{i,0})$;

- incurred losses $I_{i,j}$ are given by the backwards recursion

$$I_{i,j-1} = I_{i,j} \exp(-\zeta_{i,j-1}),$$

with initial value $I_{i,I} = P_{i,I}$.

- The components of Θ are independent and $\sigma_j, \tau_j > 0$ for all j .

7.2 Parameter estimation

Parameters Θ in the model are in general not known and need to be estimated from observations. They are estimated in a Bayesian framework. In the

Bayesian PIC model they assume that the previous assumptions hold true with deterministic $\sigma_0, \dots, \sigma_J$ and $\tau_0, \dots, \tau_{J-1}$ and

$$\Phi_m \sim N(\phi_m, s_m^2),$$

$$\Psi_n \sim N(\psi_n, t_n^2).$$

This is not a full Bayesian approach but has the advantage to give analytical expressions for the posterior distributions and the prediction uncertainty.

The Paid-incurred Chain model is implemented in the `ChainLadder` package via the function `PaidIncurredChain`. As an example we apply the function to the USAA paid and incurred triangles:

```
R> library(ChainLadder)
R> PIC <- PaidIncurredChain(USAApaid, USAAincurred)
R> PIC
```

```
$Ult.Loss.Origin
      [,1]
[1,]  983113
[2,] 1078697
[3,] 1145761
[4,] 1245171
[5,] 1371964
[6,] 1433857
[7,] 1415964
[8,] 1410065
[9,] 1320415
```

```
$Ult.Loss
[1] 11405008
```

```
$Res.Origin
      [,1]
[1,]   965.3
[2,]  3159.6
[3,]  7386.4
[4,] 18521.4
[5,] 47232.4
[6,] 113727.4
[7,] 230663.7
[8,] 443903.3
[9,] 778393.5
```

```
$Res.Tot
[1] 1643953
```



```
$s.e.  
[1] 113940
```

We can access the reserves by origin year via

```
R> PIC$Res.Origin
```

```
      [,1]  
[1,]    965.3  
[2,]   3159.6  
[3,]   7386.4  
[4,]  18521.4  
[5,]  47232.4  
[6,] 113727.4  
[7,] 230663.7  
[8,] 443903.3  
[9,] 778393.5
```

and the total reserve via

```
R> PIC$Res.Tot
```

```
[1] 1643953
```

`s.e.` is the square root of mean square error of prediction for the total ultimate loss.

It's important to notice that the model is implemented in the special case of non-informative priors for Φ_m and Ψ_n ; this means that we let $s_m^2 \rightarrow \infty$ and $t_n^2 \rightarrow \infty$.

8 One year claims development result

The stochastic claims reserving methods considered above predict the lower (unknown) triangle and assess the uncertainty of this prediction. For instance, Mack's uncertainty formula quantifies the total prediction uncertainty of the chain-ladder predictor over the entire run-off of the outstanding claims. Modern solvency considerations, such as Solvency II, require a second view of claims reserving uncertainty. This second view is a short-term view because it requires assessments of the one-year changes of the claims predictions when one updates the available information at the end of each accounting year. At time $t \geq n$ we have information

$$\mathcal{D}_t = \{C_{i,k}; i + k \leq t + 1\}.$$

This motivates the following sequence of predictors for the ultimate claim $C_{i,K}$ at times $t \geq n$

$$\hat{C}_{i,K}^{(t)} = \mathbb{E}[C_{i,K} | \mathcal{D}_t].$$

The one year claims development results (CDR), see Merz-Wüthrich [MW08a, MW14], consider the changes in these one year updates, that is,

$$\text{CDR}_{i,t+1} = \hat{C}_{i,K}^{(t)} - \hat{C}_{i,K}^{(t+1)}.$$

The tower property of conditional expectation implies that the CDRs are on average 0, that is, $\mathbb{E}[\text{CDR}_{i,t+1}|\mathcal{D}_t] = 0$ and the Merz-Wüthrich formula [MW08a, MW14] assesses the uncertainty of these predictions measured by the following conditional mean square error of prediction (MSEP)

$$\text{mse}_{\text{CDR}_{i,t+1}|\mathcal{D}_t}(0) = \mathbb{E} \left[(\text{CDR}_{i,t+1} - 0)^2 \middle| \mathcal{D}_t \right].$$

The major difficulty in the evaluation of the conditional MSEP is the quantification of parameter estimation uncertainty.

8.1 CDR functions

The one year claims development result (CDR) can be estimate via the generic CDR function for objects of `MackChainLadder` and `BootChainLadder`.

Further, the `tweedieReserve` function offers also the option to estimate the one year CDR, by setting the argument `rereserving=TRUE`.

For example, to reproduce the results of [MW14] use:

```
R> M <- MackChainLadder(MW2014, est.sigma="Mack")
R> cdrM <- CDR(M)
R> round(cdrM, 1)
```

	IBNR	CDR(1)S.E.	Mack.S.E.
1	0.0	0.0	0.0
2	1.0	0.4	0.4
3	10.1	2.5	2.6
4	21.2	16.7	16.9
5	117.7	156.4	157.3
6	223.3	137.7	207.2
7	361.8	171.2	261.9
8	469.4	70.3	292.3
9	653.5	271.6	390.6
10	1008.8	310.1	502.1
11	1011.9	103.4	486.1
12	1406.7	632.6	806.9
13	1492.9	315.0	793.9
14	1917.6	406.1	891.7
15	2458.2	285.2	916.5
16	3384.3	668.2	1106.1
17	9596.6	733.2	1295.7
Total	24134.9	1842.9	3233.7

To review the full claims development picture set the argument dev="all":

```
R> cdrAll <- CDR(M,dev="all")
R> round(cdrAll, 1)
```

	IBNR	CDR(1)S.E.	CDR(2)S.E.	CDR(3)S.E.	CDR(4)S.E.	CDR(5)S.E.	CDR(6)S.E.
1	0.0	0.0	0.0	0.0	0.0	0.0	0.0
2	1.0	0.4	0.0	0.0	0.0	0.0	0.0
3	10.1	2.5	0.4	0.0	0.0	0.0	0.0
4	21.2	16.7	2.4	0.3	0.0	0.0	0.0
5	117.7	156.4	16.4	2.4	0.3	0.0	0.0
6	223.3	137.7	154.0	16.1	2.3	0.3	0.0
7	361.8	171.2	131.0	148.0	15.5	2.2	0.3
8	469.4	70.3	185.0	141.8	160.9	16.8	2.4
9	653.5	271.6	61.8	178.0	136.9	155.7	16.3
10	1008.8	310.1	274.6	59.0	180.4	138.6	158.1
11	1011.9	103.4	293.0	260.0	53.0	170.9	131.4
12	1406.7	632.6	102.3	302.2	268.7	52.8	176.6
13	1492.9	315.0	572.1	86.6	273.0	242.8	45.4
14	1917.6	406.1	313.3	573.0	84.4	273.1	243.1
15	2458.2	285.2	395.5	305.3	560.8	80.1	267.1
16	3384.3	668.2	271.7	380.2	293.3	540.6	75.7
17	9596.6	733.2	645.4	261.0	367.0	282.8	522.9
Total	24134.9	1842.9	1485.1	1208.3	1071.1	901.1	785.3
	CDR(7)S.E.	CDR(8)S.E.	CDR(9)S.E.	CDR(10)S.E.	CDR(11)S.E.	CDR(12)S.E.	
1	0.0	0.0	0.0	0.0	0.0	0.0	0.0
2	0.0	0.0	0.0	0.0	0.0	0.0	0.0
3	0.0	0.0	0.0	0.0	0.0	0.0	0.0
4	0.0	0.0	0.0	0.0	0.0	0.0	0.0
5	0.0	0.0	0.0	0.0	0.0	0.0	0.0
6	0.0	0.0	0.0	0.0	0.0	0.0	0.0
7	0.0	0.0	0.0	0.0	0.0	0.0	0.0
8	0.3	0.0	0.0	0.0	0.0	0.0	0.0
9	2.3	0.3	0.0	0.0	0.0	0.0	0.0
10	16.6	2.4	0.3	0.0	0.0	0.0	0.0
11	150.4	15.7	2.3	0.3	0.0	0.0	0.0
12	135.8	155.6	16.3	2.3	0.3	0.0	0.0
13	159.7	122.9	141.1	14.8	2.1	0.3	0.3
14	44.1	159.9	123.0	141.4	14.8	2.1	2.1
15	237.9	42.3	156.4	120.4	138.5	14.5	14.5
16	257.3	229.3	39.9	150.8	116.1	133.6	133.6
17	71.8	248.8	221.7	38.1	145.9	112.3	112.3
Total	525.2	476.3	366.4	269.3	245.0	180.4	180.4
	CDR(13)S.E.	CDR(14)S.E.	CDR(15)S.E.	CDR(16)S.E.	CDR(17)S.E.	Mack.S.E.	
1	0.0	0.0	0.0	0.0	0	0.0	0.0
2	0.0	0.0	0.0	0.0	0	0.4	0.4
3	0.0	0.0	0.0	0.0	0	2.6	2.6

4	0.0	0.0	0.0	0.0	0	16.9
5	0.0	0.0	0.0	0.0	0	157.3
6	0.0	0.0	0.0	0.0	0	207.2
7	0.0	0.0	0.0	0.0	0	261.9
8	0.0	0.0	0.0	0.0	0	292.3
9	0.0	0.0	0.0	0.0	0	390.6
10	0.0	0.0	0.0	0.0	0	502.1
11	0.0	0.0	0.0	0.0	0	486.1
12	0.0	0.0	0.0	0.0	0	806.9
13	0.0	0.0	0.0	0.0	0	793.9
14	0.3	0.0	0.0	0.0	0	891.7
15	2.1	0.3	0.0	0.0	0	916.5
16	14.0	2.0	0.3	0.0	0	1106.1
17	129.3	13.5	1.9	0.3	0	1295.7
Total	130.1	13.7	2.0	0.3	0	3233.7

See the help files to `CDR` and `tweedieReserve` for more details.

9 Model Validation with `tweedieReserve`

Model validation is one of the key activities when an insurance company goes through the Internal Model Approval Process with the regulator. This section gives some examples how the arguments of the `tweedieReserve` function can be used to validate a stochastic reserving model. The argument `design.type` allows us to test different regression structures. The classic over-dispersed Poisson (ODP) model uses the following structure:

$$Y \sim as.factor(OY) + as.factor(DY),$$

(i.e. `design.type=c(1,1,0)`). This allows, together with the log link, to achieve the same results of the (volume weighted) chain-ladder model, thus the same model implied assumptions. A common model shortcoming is when the residuals plotted by calendar period start to show a pattern, which chain-ladder isn't capable to model. In order to overcome this, the user could be then interested to change the regression structure in order to try to strip out these patterns [GS05]. For example, a regression structure like:

$$Y \sim as.factor(DY) + as.factor(CY),$$

i.e. `design.type=c(0,1,1)` could be considered instead. This approach returns the same results of the arithmetic separation method, modelling explicitly inflation parameters between consequent calendar periods. Another interesting assumption is the assumed underlying distribution. The ODP model assumes the following:

$$P_{i,j} \sim ODP(m_{i,j}, \phi \cdot m_{i,j}),$$

which is a particular case of a Tweedie distribution, with p parameter equals to 1. Generally speaking, for any random variable Y that obeys a Tweedie distribution, the variance $\mathbb{V}[Y]$ relates to the mean $\mathbb{E}[Y]$ by the following law:

$$\mathbb{V}[Y] = a \cdot \mathbb{E}[Y]^p,$$

where a and p are positive constants. The user is able to test different p values through the `var.power` function argument. Besides, in order to validate the Tweedie's p parameter, it could be interesting to plot the likelihood profile at defined p values (through the `p.check` argument) for a given a dataset and a regression structure. This could be achieved setting the `p.optim=TRUE` argument, as follows:

```
R> p_profile <- tweedieReserve(MW2008, p.optim=TRUE,
  p.check=c(0,1.1,1.2,1.3,1.4,1.5,2,3),
  design.type=c(0,1,1),
  rereserving=FALSE,
  bootstrap=0,
  progressBar=FALSE)
R> # 0 1.1 1.2 1.3 1.4 1.5 2 3
R> # .....Done.
R> # MLE of p is between 0 and 1, which is impossible.
R> # Instead, the MLE of p has been set to NA .
R> # Please check your data and the call to tweedie.profile().
R> # Error in if ((xi.max == xi.vec[1]) | (xi.max == xi.vec[length(xi.vec)])) { :
R> # missing value where TRUE/FALSE needed
```

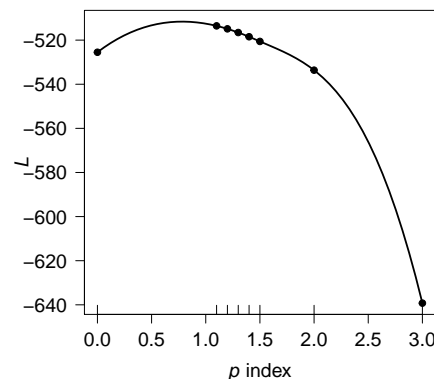


Figure 7: Likelihood profile of regression structure

This example shows, see Figure 7, that the MLE of p seems to be between 0 and 1, which is not possible as Tweedie models aren't defined for $0 < p$

< 1 , thus the Error message. But, despite this, we can conclude that overall a value $p=1$ could be reasonable for this dataset and the chosen regression function, as anyway it seems to be near the MLE. Other sensitivities could be run on:

- Bootstrap type (parametric / semi-parametric), via the `bootstrap` argument
- Bias adjustment (if using semi-parametric bootstrap), via the `boot.adj` argument

Please refer to `help(tweedieReserve)` for additional information.

10 Using ChainLadder with RExcel and SWord

The ChainLadder package comes with example files which demonstrate how its functions can be embedded in Excel and Word using the [statconn](#) interface [\[BN07\]](#).

The spreadsheet is located in the Excel folder of the package. The R command

```
R> system.file("Excel", package="ChainLadder")
```

will tell you the exact path to the directory. To use the spreadsheet you will need the RExcel-Add-in [\[BN07\]](#). The package also provides an example SWord file, demonstrating how the functions of the package can be integrated into a MS Word file via SWord [\[BN07\]](#). Again you find the Word file via the command:

```
R> system.file("SWord", package="ChainLadder")
```

The package comes with several demos to provide you with an overview of the package functionality, see

```
R> demo(package="ChainLadder")
```

11 Further resources

Other useful documents and resources to get started with R in the context of actuarial work:

- Introduction to R for Actuaries [\[DS06\]](#).
- Computational Actuarial Science with R [\[Cha14\]](#)
- Modern Actuarial Risk Theory – Using R [\[KGDD01\]](#)
- An Actuarial Toolkit [\[MSH⁺06\]](#).
- Mailing list [R-SIG-insurance](#)⁸: Special Interest Group on using R in actuarial science and insurance

⁸<https://stat.ethz.ch/mailman/listinfo/r-sig-insurance>

11.1 Other insurance related R packages

Below is a list of further R packages in the context of insurance. The list is by no-means complete, and the CRAN Task Views '[Empirical Finance](#)' and '[Probability Distributions](#)' will provide links to additional resources. Please feel free to contact [us](#) with items to be added to the list.

- `cp1m`: Likelihood-based and Bayesian methods for fitting Tweedie compound Poisson linear models [[Zha12](#)].
- `lossDev`: A Bayesian time series loss development model. Features include skewed-t distribution with time-varying scale parameter, Reversible Jump MCMC for determining the functional form of the consumption path, and a structural break in this path [[LS11](#)].
- `favir`: Formatted Actuarial Vignettes in R. FAViR lowers the learning curve of the R environment. It is a series of peer-reviewed Sweave papers that use a consistent style [[Esc11](#)].
- `actuar`: Loss distributions modelling, risk theory (including ruin theory), simulation of compound hierarchical models and credibility theory [[DGP08](#)].
- `fitdistrplus`: Help to fit of a parametric distribution to non-censored or censored data [[DMPDD10](#)].
- `mondate`: R package to keep track of dates in terms of months [[Mur11](#)].
- `lifecontingencies`: Package to perform actuarial evaluation of life contingencies [[Spe11](#)].
- `MRMR`: Multivariate Regression Models for Reserving [[Fan13](#)].

11.2 Presentations

Over the years the contributors of the ChainLadder package have given numerous presentations and most of those are still available on-line:

- A re-reserving algorithm to derive the one-year reserve risk view, R in Insurance, London, Alessandro Carrato, 2013.
- Bayesian Hierarchical Models in Property-Casualty Insurance, Wayne Zhang, 2011
- [ChainLadder at the Predictive Modelling Seminar, Institute of Actuaries, November 2010](#), Markus Gesmann, 2011
- [Reserve variability calculations](#), CAS spring meeting, San Diego, Jimmy Curcio Jr., Markus Gesmann and Wayne Zhang, 2010
- [The ChainLadder package, working with databases and MS Office interfaces, presentation at the "R you ready?" workshop](#), Institute of Actuaries, Markus Gesmann, 2009
- [The ChainLadder package](#), London R user group meeting, Markus Gesmann, 2009

- **Introduction to R, Loss Reserving with R**, Stochastic Reserving and Modelling Seminar, Institute of Actuaries, Markus Gesmann, 2008
- **Loss Reserving with R**, CAS meeting, Vincent Goulet, Markus Gesmann and Daniel Murphy, 2008
- **The ChainLadder package** R-user conference Dortmund, Markus Gesmann, 2008

11.3 Further reading

Other papers and presentations which cited ChainLadder : [Orr07], [Nic09], [Zha10], [MNNV10], [Sch10], [MNV10], [Esc11], [Spe11]

12 Training and consultancy

Please contact **us** if you would like to discuss tailored training or consultancy.

References

- [BBMW06] M. Buchwalder, H. Bühlmann, M. Merz, and M.V. Wüthrich. The mean square error of prediction in the chain ladder reserving method (mack and murphy revisited). *North American Actuarial Journal*, 36:521–542, 2006.
- [BN07] Thomas Baier and Erich Neuwirth. Excel :: Com :: R. *Computational Statistics*, 22(1), April 2007. Physica Verlag.
- [Cha14] Arthur Charpentier, editor. *Computational Actuarial Science with R*. Chapman and Hall/CRC, 2014.
- [Cla03] David R. Clark. *LDF Curve-Fitting and Stochastic Reserving: A Maximum Likelihood Approach*. Casualty Actuarial Society, 2003. CAS Fall Forum.
- [DGP08] C Dutang, V. Goulet, and M. Pigeon. actuar: An R package for actuarial science. *Journal of Statistical Software*, 25(7), 2008.
- [DMPDD10] Marie Laure Delignette-Muller, Regis Pouillot, Jean-Baptiste Denis, and Christophe Dutang. *fitdistrplus: help to fit of a parametric distribution to non-censored or censored data*, 2010. R package version 0.1-3.
- [DS06] Nigel De Silva. An introduction to r: Examples for actuaries. <http://toolkit.pbwiki.com/RToolkit>, 2006.
- [Esc11] Benedict Escoto. *favir: Formatted Actuarial Vignettes in R*, 0.5-1 edition, January 2011.
- [Fan13] Brian A. Fannin. *MRMR: Multivariate Regression Models for Reserving*, 2013. R package version 0.1.3.

- [GBB⁺09] Brian Gravelsons, Matthew Ball, Dan Beard, Robert Brooks, Naomi Couchman, Brian Gravelsons, Charlie Kefford, Darren Michaels, Patrick Nolan, Gregory Overton, Stephen Robertson-Dunn, Emiliano Ruffini, Graham Sandhouse, Jerome Schilling, Dan Sykes, Peter Taylor, Andy Whiting, Matthew Wilde, and John Wilson. B12: Uk asbestos working party update 2009. <http://www.actuaries.org.uk/research-and-resources/documents/b12-uk-asbestos-working-party-update-2009-5mb>, October 2009. Presented at the General Insurance Convention.
- [Ges14] Markus Gesmann. Claims reserving and IBNR. In *Computational Actuarial Science with R*, pages 656–Page. Chapman and Hall/CRC, 2014.
- [GMZ⁺15] Markus Gesmann, Dan Murphy, Wayne Zhang, Alessandro Carato, Mario Wüthrich, and Fabio Concina. *ChainLadder: Statistical Methods and Models for Claims Reserving in General Insurance*, 2015. R package version 0.2.1.
- [GS05] Gigante and Sigalotti. Model risk in claims reserving with glm. *Giornale dell IIA*, LXVIII:55 – 87, 2005.
- [KGDD01] R. Kaas, M. Goovaerts, J. Dhaene, and M. Denuit. *Modern actuarial risk theory*. Kluwer Academic Publishers, Dordrecht, 2001.
- [LFK⁺02] Graham Lyons, Will Forster, Paul Kedney, Ryan Warren, and Helen Wilkinson. *Claims Reserving Working Party paper*. Institute of Actuaries, October 2002.
- [LS11] Christopher W. Laws and Frank A. Schmid. *lossDev: Robust Loss Development Using MCMC*, 2011. R package version 3.0.0-1.
- [Mac93] Thomas Mack. Distribution-free calculation of the standard error of chain ladder reserve estimates. *ASTIN Bulletin*, 23:213–225, 1993.
- [Mac99] Thomas Mack. The standard error of chain ladder reserve estimates: Recursive calculation and inclusion of a tail factor. *Astin Bulletin*, Vol. 29(2):361 – 266, 1999.
- [Mic02] Darren Michaels. APH: how the love carnal and silicone implants nearly destroyed Lloyd’s (slides). <http://www.actuaries.org.uk/research-and-resources/documents/aph-how-love-carnal-and-silicone-implants-nearly-destroyed-lloyds-s>, December 2002. Presented at the Younger Members’ Convention.
- [MNNV10] Maria Dolores Martinez Miranda, Bent Nielsen, Jens Perch Nielsen, and Richard Verrall. *Cash flow simulation for a model of outstanding liabilities based on claim amounts and claim numbers*. CASS, September 2010.

- [MNV10] Maria Dolores Martinez Miranda, Jens Perch Nielsen, and Richard Verrall. *Double Chain Ladder*. ASTIN, Colloquia Madrid edition, 2010.
- [MSH⁺06] Trevor Maynard, Nigel De Silva, Richard Holloway, Markus Gesmann, Sie Lau, and John Harnett. An actuarial toolkit. introducing The Toolkit Manifesto. <http://www.actuaries.org.uk/sites/all/files/documents/pdf/actuarial-toolkit.pdf>, 2006. General Insurance Convention.
- [Mur94] Daniel Murphy. Unbiased loss development factors. *PCAS*, 81:154 – 222, 1994.
- [Mur11] Daniel Murphy. *mondate: Keep track of dates in terms of months*, 2011. R package version 0.9.8.24.
- [MW08a] Michael Merz and Mario V. Wüthrich. Modelling the claims development result for solvency purposes. *CAS E-Forum*, Fall:542–568, 2008.
- [MW08b] Michael Merz and Mario V. Wüthrich. Prediction error of the multivariate chain ladder reserving method. *North American Actuarial Journal*, 12:175–197, 2008.
- [MW10] M. Merz and M. Wüthrich. Paid-incurred chain claims reserving method. *Insurance: Mathematics and Economics*, 46(3):568–579, 2010.
- [MW14] Michael Merz and Mario V. Wüthrich. Claims run-off uncertainty: the full picture. *SSRN Manuscript*, 2524352, 2014.
- [Nic09] Luke Nichols. *Multimodel Inference for Reserving*. Australian Prudential Regulation Authority (APRA), December 2009.
- [Orr07] James Orr. *A Simple Multi-State Reserving Model*. ASTIN, Colloquia Orlando edition, 2007.
- [Orr12] James Orr. *GIROC reserving research workstream*. Institute of Actuaries, November 2012.
- [PR02] P.D.England and R.J.Verrall. Stochastic claims reserving in general insurance. *British Actuarial Journal*, 8:443–544, 2002.
- [PS05] Carsten Pröhl and Klaus D. Schmidt. Multivariate chain-ladder. *Dresdner Schriften zur Versicherungsmathematik*, 2005.
- [QM04] Gerhard Quarg and Thomas Mack. Munich chain ladder. Munich Re Group, 2004.
- [Sch10] Ernesto Schirmacher. Reserve variability calculations, chain ladder, R, and Excel. <http://www.casact.org/affiliates/cane/0910/schirmacher.pdf>, September 2010. Presentation at the Casualty Actuaries of New England (CANE) meeting.
- [Sch11] Klaus D. Schmidt. A bibliography on loss reserving. <http://www.math.tu-dresden.de/sto/schmidt/dsvm/reserve.pdf>, 2011.

- [Spe11] Giorgio Alfredo Spedicato. *Introduction to lifecontingencies Package*. StatisticalAdvisor Inc, 0.0.4 edition, November 2011.
- [Tea12a] R Development Core Team. *R Data Import/Export*. R Foundation for Statistical Computing, 2012. ISBN 3-900051-10-0.
- [Tea12b] R Development Core Team. *R Installation and Administration*. R Foundation for Statistical Computing, 2012. ISBN 3-900051-09-7.
- [ZB00] Ben Zehnwrith and Glen Barnett. Best estimates for reserves. *Proceedings of the CAS*, LXXXVII(167), November 2000.
- [ZDG12] Yanwei Zhang, Vanja Dukic, and James Guszcza. A bayesian nonlinear model for forecasting insurance loss payments. *Journal of the Royal Statistical Society, Series A*, 175:637–656, 2012.
- [Zha10] Yanwei Zhang. A general multivariate chain ladder model. *Insurance: Mathematics and Economics*, 46:588 – 599, 2010.
- [Zha12] Yanwei Zhang. Likelihood-based and bayesian methods for tweedie compound poisson linear mixed models. *Statistics and Computing*, 2012. forthcoming.